

```

1 . set linesize 80

2 .
3 . *****
> ****
4 . ***** We summarize the main effects models that were supposed to be signif
> icant
5 . *****
> ****
6 . set linesize 80

7 . est clear

8 .
9 .
10 . * work
11 . forvalues j=2/2 {
    2. local w2bf bf1 bf4 bf2 bf4m bf5m bf7m bf8 bf15m bf17 bf20 bf22 bf29 bf30
    > bf40
    3. set more off
    4. di as input "For females trimmed HP2work on wave 2 with dose ns"
    5.      des age occ1w`j' -occ8w`j' inclw`j' -inc4w`j' avgcumdosew`j' bf8 /
    > //
    >      marrw`j'1 -marrw`j'6 `w2bf'
    6.      logistic HP2work age whppain whpsleep whpel ///
    >      avgcumdosew2 bf8 illw`j' shjobw`j' radhlw`j' if gender==1, c
    > oef nolog
    7.
12 . eststo work
    8.      }
    For females trimmed HP2work on wave 2 with dose ns

```

variable name	storage type	display format	value label	variable label
age	double	%8.0g		* Respondent's age
occ1w2	double	%15.0g	LABJ	profess executive administration in 1996
occ2w2	double	%15.0g	LABJ	technical sales admin support in 1996
occ3w2	double	%15.0g	LABJ	service occup protective services in 1996
occ4w2	double	%15.0g	LABJ	precision prod mechan craft construction in 1996
occ5w2	double	%15.0g	LABJ	factory laborer machinist transp cleaner in 1996
occ6w2	double	%15.0g	LABJ	farming agricul forestry fishing trapping logging in 1996
occ7w2	double	%15.0g	LABJ	homemaking caregiving in 1996

occ8w2	double	%15.0g	LABJ	student in 1996
inclw2	double	%15.0g	LABJ	Income is not sufficient for basic neccessities in 1996
inc2w2	double	%15.0g	LABJ	Income is just sufficient for basic neccessities in 1996
inc3w2	double	%15.0g	LABJ	Income is sufficient for basics plus extra purchases/savings in 1996
inc4w2	double	%15.0g	LABJ	Income allows to comfortably afford luxury items in 1996
avgcumdosew2	double	%8.0g		Average mean dose CS1337 in mGy for wave 2
bf8	float	%9.0g		bf8 = max(0, radtlw3 - 40) * bf5m
marrw21	byte	%8.0g		marrw2==1. single
marrw22	byte	%8.0g		marrw2==2. cohabitating
marrw23	byte	%8.0g		marrw2==3. married
marrw24	byte	%8.0g		marrw2==4. separated
marrw25	byte	%8.0g		marrw2==5. divorced
marrw26	byte	%8.0g		marrw2==6. widowed
bf1	float	%9.0g		bf1 = max(0, kzchorn - 40)
bf4	float	%9.0g		bf4 = max(0, 24 - BSIsoma)
bf2	float	%9.0g		bf2 = max(0, efradw3 - 1.61252E-006) * bf1
bf4m	float	%9.0g		bf4m = max(0, 32 - BSIsoma)
bf5m	float	%9.0g		bf5m = max(0, ecprw3 - 75) * bf4m
bf7m	float	%9.0g		bf7m = max(0, radtlw3 - 2.12558E-007) * bf4m
bf8	float	%9.0g		bf8 = max(0, radtlw3 - 40) * bf5m
bf15m	float	%9.0g		bf15m= max(0, 1 - icdxcnt) * bf2
bf17	float	%9.0g		bf17 = max(0, 20 - ecprw3) * bf15m
bf20	float	%9.0g		bf20 = max(0, kzchorn - 2.53946E-006)
bf22	float	%9.0g		bf22 = max(0, icdxcnt - 1.01635E-007) * bf7m
bf29	float	%9.0g		bf29 = max(0, 18 - CSsocspt) * bf8
bf30	float	%9.0g		bf30 = max(0, neiwl - 85) * bf20
bf40	float	%9.0g		bf40 = max(0, icdxcnt - 1.01635E-007)

Logistic regression

Log likelihood = **-139.93087**

Number of obs	=	340
LR chi2(9)	=	65.88
Prob > chi2	=	0.0000
Pseudo R2	=	0.1906

HP2work	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0133896	.0138973	0.96	0.335	-.0138487	.0406279
whppain	.0277922	.0099062	2.81	0.005	.0083765	.047208
whpsleep	.0022041	.0066476	0.33	0.740	-.0108249	.0152332
whpel	.0167666	.005548	3.02	0.003	.0058927	.0276405
avgcumdosew2	-.0156145	.049985	-0.31	0.755	-.1135833	.0823542
bf8	-.0000491	.0000255	-1.93	0.054	-.0000991	8.70e-07
illw2	.3746153	.238162	1.57	0.116	-.0921737	.8414043
shjobw2	.0034013	.0040952	0.83	0.406	-.0046252	.0114277
radhlw2	.0068395	.004953	1.38	0.167	-.0028683	.0165472
_cons	-3.455413	.727987	-4.75	0.000	-4.882241	-2.028584

```

13 .
14 .
15 .
16 .
17 . * home care
18 .
19 . * Dose work relationship for females in wave 2 washes out also
20 . set more off

21 . forvalues j=2/2 {
    2. di _skip(4)
    3. di as input "For females hp2hmcare on wave 1 with dose ns"
    4. des age avgcumdosew`j' `w2bf'
    5. logistic HP2work age havmilsq radhlw2 whppain whpsleep whpel ///
> avgcumdosew2 shhlw`j' shjobw`j' shrelaw`j' suprtw`j' if gender==2, coef n
> olog
    6. eststo hmcare
    7. }

```

For females hp2hmcare on wave 1 with dose ns

variable name	storage type	display format	value label	variable label
age	double	%8.0g		* Respondent's age
avgcumdosew2	double	%8.0g		Average mean dose CS1337 in mGy for wave 2
bf1	float	%9.0g		bf1 = max(0, kzchorn - 40)
bf4	float	%9.0g		bf4 = max(0, 24 - BSIsoma)
bf2	float	%9.0g		bf2 = max(0, efradw3 - 1.61252E-006) * bf1
bf4m	float	%9.0g		bf4m = max(0, 32 - BSIsoma)
bf5m	float	%9.0g		bf5m = max(0, ecprw3 - 75) * bf4m

```

bf7m          float   %9.0g          bf7m = max(0, radtlw3 -
                                         2.12558E-007) * bf4m
bf8           float   %9.0g          bf8 = max(0, radtlw3 - 40) *
                                         bf5m
bf15m         float   %9.0g          bf15m= max(0, 1 - icdxcnt) * bf2
bf17          float   %9.0g          bf17 = max(0, 20 - ecprw3) *
                                         bf15m
bf20          float   %9.0g          bf20 = max(0, kzchorn -
                                         2.53946E-006)
bf22          float   %9.0g          bf22 = max(0, icdxcnt -
                                         1.01635E-007) * bf7m
bf29          float   %9.0g          bf29 = max(0, 18 - CSsocspt) *
                                         bf8
bf30          float   %9.0g          bf30 = max(0, neiwl - 85) * bf20
bf40          float   %9.0g          bf40 = max(0, icdxcnt -
                                         1.01635E-007)

```

Logistic regression

```

Number of obs   =      363
LR chi2(11)     =      87.33
Prob > chi2     =      0.0000
Pseudo R2      =      0.2114

```

Log likelihood = **-162.89904**

HP2work	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0366858	.0131691	2.79	0.005	.0108748	.0624969
havmilsq	-4.12e-07	8.51e-07	-0.48	0.628	-2.08e-06	1.25e-06
radhlw2	.0079654	.0043944	1.81	0.070	-.0006474	.0165782
whppain	.010523	.0076361	1.38	0.168	-.0044435	.0254895
whpsleep	.0108514	.0052697	2.06	0.039	.0005229	.0211799
whpel	.0098934	.0048783	2.03	0.043	.0003321	.0194546
avgcumdosew2	.0787768	.0869756	0.91	0.365	-.0916923	.2492458
shhlw2	.0060954	.0050051	1.22	0.223	-.0037144	.0159052
shjobw2	-.0029385	.004756	-0.62	0.537	-.0122602	.0063831
shrelaw2	-.0125997	.0051163	-2.46	0.014	-.0226275	-.0025719
suprtw2	.0008366	.0034693	0.24	0.809	-.0059631	.0076363
_cons	-4.409484	.7893441	-5.59	0.000	-5.95657	-2.862398

```

22 .
23 .
24 . * social problems
25 . forvalues j=2/2 {
      2. set more off
      3. di as input "females trimmed HP2probsoc on wave 2 with dose ns"
      4.      des age occ1w`j'-occ8w`j' inc1w`j'-inc4w`j' avgcumdosew`j' bf8 /
> //
>      marrw`j'1-marrw`j'6 `w2bf'
      5.      logistic HP2probsoc age avgcumdosew2 whppain whpsleep whpel ///
>      bf8 illw`j' shjobw`j' radhlw`j' if gender==1, coef nolog
>
      6. eststo socprobs
      7.      }
females trimmed HP2probsoc on wave 2 with dose ns

```

variable name	storage type	display format	value label	variable label
age	double	%8.0g		* Respondent's age
occ1w2	double	%15.0g	LABJ	profess executive administration in 1996
occ2w2	double	%15.0g	LABJ	technical sales admin support in 1996
occ3w2	double	%15.0g	LABJ	service occup protective services in 1996
occ4w2	double	%15.0g	LABJ	precision prod mechan craft construction in 1996
occ5w2	double	%15.0g	LABJ	factory laborer machinist transp cleaner in 1996
occ6w2	double	%15.0g	LABJ	farming agricul forestry fishing trapping logging in 1996
occ7w2	double	%15.0g	LABJ	homemaking caregiving in 1996
occ8w2	double	%15.0g	LABJ	student in 1996
inc1w2	double	%15.0g	LABJ	Income is not sufficient for basic neccessities in 1996
inc2w2	double	%15.0g	LABJ	Income is just sufficient for basic neccessities in 1996
inc3w2	double	%15.0g	LABJ	Income is sufficient for basics plus extra purchases/savings in 1996
inc4w2	double	%15.0g	LABJ	Income allows to comfortably afford luxury items in 1996
avgcumdosew2	double	%8.0g		Average mean dose CS1337 in mGy for wave 2
bf8	float	%9.0g		bf8 = max(0, radtlw3 - 40) * bf5m
marrw21	byte	%8.0g		marrw2==1. single
marrw22	byte	%8.0g		marrw2==2. cohabitating

```

marrw23      byte    %8.0g      marrw2==3. married
marrw24      byte    %8.0g      marrw2==4. separated
marrw25      byte    %8.0g      marrw2==5. divorced
marrw26      byte    %8.0g      marrw2==6. widowed
bf1          float   %9.0g      bf1 = max(0, kzchorn - 40)
bf4          float   %9.0g      bf4 = max(0, 24 - BSIsoma)
bf2          float   %9.0g      bf2 = max(0, efradw3 -
                                1.61252E-006) * bf1
bf4m         float   %9.0g      bf4m = max(0, 32 - BSIsoma)
bf5m         float   %9.0g      bf5m = max(0, ecprw3 - 75) *
                                bf4m
bf7m         float   %9.0g      bf7m = max(0, radtlw3 -
                                2.12558E-007) * bf4m
bf8          float   %9.0g      bf8 = max(0, radtlw3 - 40) *
                                bf5m
bf15m        float   %9.0g      bf15m= max(0, 1 - icdxcnt) * bf2
bf17         float   %9.0g      bf17 = max(0, 20 - ecprw3) *
                                bf15m
bf20         float   %9.0g      bf20 = max(0, kzchorn -
                                2.53946E-006)
bf22         float   %9.0g      bf22 = max(0, icdxcnt -
                                1.01635E-007) * bf7m
bf29         float   %9.0g      bf29 = max(0, 18 - CSsocspt) *
                                bf8
bf30         float   %9.0g      bf30 = max(0, neiwl - 85) * bf20
bf40         float   %9.0g      bf40 = max(0, icdxcnt -
                                1.01635E-007)

```

Logistic regression

```

Number of obs    =      340
LR chi2(9)       =      91.72
Prob > chi2      =      0.0000
Pseudo R2       =      0.3664

```

Log likelihood = **-79.291415**

HP2probsoc	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.037799	.0194419	1.94	0.052	-.0003064	.0759044
avgcumdosew2	.0083805	.0536714	0.16	0.876	-.0968135	.1135745
whppain	.0338775	.012604	2.69	0.007	.0091742	.0585808
whpsleep	.0053094	.0083267	0.64	0.524	-.0110106	.0216293
whpel	.0199609	.0076543	2.61	0.009	.0049587	.034963
bf8	-.0000415	.0000304	-1.37	0.172	-.0001011	.000018
illw2	.3212823	.3100339	1.04	0.300	-.286373	.9289375
shjobw2	.0121068	.0059195	2.05	0.041	.0005047	.0237088
radhlw2	.0179923	.0070408	2.56	0.011	.0041926	.0317919
_cons	-7.244506	1.207454	-6.00	0.000	-9.611072	-4.87794

```
*****  
> *  
*****  
  
> *  
*****  
  
> *  
*****  
  
***** Trimmed Female Famly problems model *****  
  
> *  
*****  
  
> *  
*****  
  
> *  
*****  
  
***** 2 Jul 2012 01:02:33 *****  
  
> *  
*****  
  
> *
```

```

*****
> *
*****
> *
*****
> *
*****
> *
*****
More partly female Trimmed wave 2
> *
*****
Dose => Problems with Family at home models
> *
*****
> *
*****

```

```

*****
> *
*****
2 Jul 2012    01:02:33    *****
> *
*****
> *
*****
> *

```

```
33 . local w2bf bf1 bf4 bf2 bf8 bf15m bf17 bf20 bf22 bf29 bf30 bf40
```

```
34 . logit HP2pbfhm age bf4 bf40 avgcumdosew2 radhlw2 ///
> whppain whpsleep whpel if gender==2, iterate(50) nolog
```

```

Logistic regression               Number of obs   =       363
                                LR chi2(8)        =       83.91
                                Prob > chi2        =       0.0000
Log likelihood = -97.941266       Pseudo R2     =       0.2999

```

HP2pbfhm	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0427925	.018401	2.33	0.020	.0067272	.0788577
bf4	-.0828459	.0504609	-1.64	0.101	-.1817474	.0160557
bf40	-.2093721	.1016708	-2.06	0.039	-.4086432	-.010101
avgcumdosew2	.0018638	.1036398	0.02	0.986	-.2012664	.2049941
radhlw2	.0096602	.0061285	1.58	0.115	-.0023514	.0216718
whppain	.0141081	.0098379	1.43	0.152	-.0051738	.03339
whpsleep	.0066821	.0064638	1.03	0.301	-.0059867	.0193508
whpel	.0146091	.0067182	2.17	0.030	.0014418	.0277765
_cons	-4.841001	1.282889	-3.77	0.000	-7.355417	-2.326586

```
35 . eststo familyprbs
```



```
44 .      eststo sexlife
```

```
7. }
```

```
trimmed HP2sexlife main effects models wave 1 for H1 part 2 with dose ns
wave 2 male dose avgcumdosew2 main effect not signif
```

```
Iteration 0:  log likelihood = -171.51396
Iteration 1:  log likelihood = -118.4729
Iteration 2:  log likelihood = -111.80137
Iteration 3:  log likelihood = -111.62448
Iteration 4:  log likelihood = -111.62389
Iteration 5:  log likelihood = -111.62389
```

```
Logistic regression                                Number of obs   =       340
                                                    LR chi2(8)      =      119.78
                                                    Prob > chi2     =       0.0000
Log likelihood = -111.62389                        Pseudo R2      =       0.3492
```

HP2sxlife	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0633624	.0158369	4.00	0.000	.0323227	.094402
bf4	-.1290276	.039812	-3.24	0.001	-.2070577	-.0509974
illw2	.2494849	.2532825	0.99	0.325	-.2469397	.7459094
whppain	.0165904	.0109372	1.52	0.129	-.0048462	.038027
whpsleep	.0050021	.0072864	0.69	0.492	-.009279	.0192832
whpel	.0072083	.0063787	1.13	0.258	-.0052937	.0197104
avgcumdosew2	.0392164	.0479775	0.82	0.414	-.0548178	.1332506
radhlw2	.0096677	.0056501	1.71	0.087	-.0014063	.0207417
_cons	-4.55453	1.106189	-4.12	0.000	-6.722621	-2.38644

```
45 .
```

```
46 . * inthob
```

```
47 .
```

```
48 . forvalues j=2/2 {
```

```
2. set more off
```

```
3.
```

```

49 . des age educ1-educ7 marrw`j'1-marrw`j'6 inclw`j'-inc4w`j' ///
    >   bf1 bf4 bf9 bf11 bf4m bf15m bf30 bf40
    4.
50 .   foreach var in HP2inthob {
    5.       forvalues k=2/2 {
    6.   local w2bf bf1 bf4 bf2 bf4m bf5m bf7m bf8 bf15m bf17 bf20 bf22 bf29 bf30
    > bf40
    7.
51 . di as input "trimmed main model for `var' for wave= `j' "
    8. di _skip(4)
    9. di as input  "chunk 8 H1 test:Gender= `k'  model Wave = `j' for `e(depva
    > r)' "
    10. di _skip(4)
    11. title "Full Nottingham Part 2 subscale models for male and then females"
    12. di as input "Model for gender==`k' and wave == `j'"
    13. di _skip(2)
    14.   logit `var' age    ///
    >           avgcumdosew`j' bf4 bf30    ///
    >           shrelaw`j' suprtw`j' whppain whpsleep whpel    ///
    >           radhlw2 if gender==2, difficult iterate(50) nolog
    15.
52 .           eststo inthobbies
    16. }
    17. }
    18. }

```

variable name	storage type	display format	value label	variable label
age	double	%8.0g		* Respondent's age
educ1	byte	%8.0g		educ==1. did not graduate high school
educ2	byte	%8.0g		educ==2. graduated high school
educ3	byte	%8.0g		educ==3. technical degree
educ4	byte	%8.0g		educ==4. did not finish college/bachelor's
educ5	byte	%8.0g		educ==5. graduated college/bachelor's
educ6	byte	%8.0g		educ==6. finished specialist/master's degree
educ7	byte	%8.0g		educ==7. doctor of science/phd
marrw21	byte	%8.0g		marrw2==1. single
marrw22	byte	%8.0g		marrw2==2. cohabitating
marrw23	byte	%8.0g		marrw2==3. married
marrw24	byte	%8.0g		marrw2==4. separated
marrw25	byte	%8.0g		marrw2==5. divorced
marrw26	byte	%8.0g		marrw2==6. widowed
inclw2	double	%15.0g	LABJ	Income is not sufficient for basic neccessities in 1996

inc2w2	double	%15.0g	LABJ	Income is just sufficient for basic neccessities in 1996
inc3w2	double	%15.0g	LABJ	Income is sufficient for basics plus extra purchases/savings in 1996
inc4w2	double	%15.0g	LABJ	Income allows to comfortably afford luxury items in 1996
bf1	float	%9.0g		bf1 = max(0, kzchorn - 40)
bf4	float	%9.0g		bf4 = max(0, 24 - BSIsoma)
bf9	float	%9.0g		bf9= max(0, 30 - shhlw1)
bf11	float	%9.0g		bf11= max(0, 20 - sufamw1)
bf4m	float	%9.0g		bf4m = max(0, 32 - BSIsoma)
bf15m	float	%9.0g		bf15m= max(0, 1 - icdxcnt) * bf2
bf30	float	%9.0g		bf30 = max(0, neiwl - 85) * bf20
bf40	float	%9.0g		bf40 = max(0, icdxcnt - 1.01635E-007)

trimmed main model for HP2inthob for wave= 2

chunk 8 H1 test:Gender= 2 model Wave = 2 for HP2sxlife

```

*****
> *
*****
> *
*****
> *
*****
> *
***** Full Nottingham Part 2 subscale models for male and then females *****
> *
*****
> *
*****
> *
*****
> *
*****
> *
*****
> *
*****
> *
*****
> *
*****
> *
*****
> *
*****

```

Model for gender==2 and wave == 2

Logistic regression	Number of obs	=	363
	LR chi2(10)	=	131.62
	Prob > chi2	=	0.0000
Log likelihood = -106.3048	Pseudo R2	=	0.3824

HP2inthob	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0576034	.0164784	3.50	0.000	.0253064	.0899004
avgcumdosew2	.0687465	.0893617	0.77	0.442	-.1063992	.2438921
bf4	-.0162435	.0495227	-0.33	0.743	-.1133062	.0808192
bf30	.000564	.0002877	1.96	0.050	2.23e-07	.0011278
shrelaw2	-.0168665	.0063941	-2.64	0.008	-.0293987	-.0043342
suprtw2	-.0130054	.004409	-2.95	0.003	-.0216469	-.0043639
whppain	.0043281	.0104002	0.42	0.677	-.0160559	.0247121
whpsleep	.0274628	.0064548	4.25	0.000	.0148117	.0401139
whpel	.0088553	.006562	1.35	0.177	-.0040059	.0217165
radhlw2	.0100719	.0059702	1.69	0.092	-.0016294	.0217732
_cons	-5.973518	1.343872	-4.45	0.000	-8.607459	-3.339576

```

53 .
54 .
55 .
56 . * vacation plans
57 .
58 . di _skip(1)

59 . di as input "For females trimmed vacation plans model on wave2 and d2 is not
    > signif "
    For females trimmed vacation plans model on wave2 and d2 is not signif

60 . di _skip(1)

61 . logistic hp2vacatn age deaw2 shjobw2 bf7m havmil ///
    > radhlw2 avgcumdosew2 whppain whpsleep whpel if ///
    > gender==2, coef nolog

Logistic regression                                Number of obs   =       363
                                                    LR chi2(10)    =       92.19
                                                    Prob > chi2    =       0.0000
Log likelihood = -121.41941                      Pseudo R2     =       0.2752

```

hp2vacatn	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.055949	.0165753	3.38	0.001	.0234619	.088436
deaw2	.1334005	.1797698	0.74	0.458	-.2189417	.4857428
shjobw2	-.0076256	.0044031	-1.73	0.083	-.0162555	.0010042
bf7m	-.000223	.0002857	-0.78	0.435	-.0007829	.0003369
havmil	-.0001227	.0010664	-0.12	0.908	-.0022129	.0019674
radhlw2	.0168136	.0060081	2.80	0.005	.0050379	.0285893
avgcumdosew2	.0422269	.0909077	0.46	0.642	-.1359489	.2204026
whppain	.0100323	.0089085	1.13	0.260	-.0074281	.0274926
whpsleep	.0120968	.0059282	2.04	0.041	.0004777	.0237159
whpel	.0069496	.0057496	1.21	0.227	-.0043194	.0182185
_cons	-6.120044	1.04653	-5.85	0.000	-8.171205	-4.068883

```
62 .     eststo vacations
```

```
63 .
```

```
64 .
```

```
65 . estout work hmcare socprobs familyprbs,      cells(b(star fmt(%9.3f)) se(
> par) t p)          ///
>          stats(r2_p bic N, fmt(%9.3f %9.0g) )      ///
>          legend collabels(none) varlabels(_cons Constant)  ///
>          mlabels(work hmcare socprobs familyprbs ) ///
>          title("main effects regression coefficients"
> ///
>          "of reconstructed dose for females in wave 2")
```

```
main effects regression coefficients of reconstructed dose for females in wave
> 2
```

	work	hmcare	socprobs	familyprbs
main				
age	0.013	0.037**	0.038	0.043*
	(0.014)	(0.013)	(0.019)	(0.018)
	0.963	2.786	1.944	2.326
	0.335	0.005	0.052	0.020
whppain	0.028**	0.011	0.034**	0.014
	(0.010)	(0.008)	(0.013)	(0.010)
	2.806	1.378	2.688	1.434
	0.005	0.168	0.007	0.152
whpsleep	0.002	0.011*	0.005	0.007
	(0.007)	(0.005)	(0.008)	(0.006)
	0.332	2.059	0.638	1.034
	0.740	0.039	0.524	0.301
whpel	0.017**	0.010*	0.020**	0.015*
	(0.006)	(0.005)	(0.008)	(0.007)

	3.022	2.028	2.608	2.175
	0.003	0.043	0.009	0.030
avgcumdosew2	-0.016	0.079	0.008	0.002
	(0.050)	(0.087)	(0.054)	(0.104)
	-0.312	0.906	0.156	0.018
	0.755	0.365	0.876	0.986
bf8	-0.000		-0.000	
	(0.000)		(0.000)	
	-1.926		-1.367	
	0.054		0.172	
illw2	0.375		0.321	
	(0.238)		(0.310)	
	1.573		1.036	
	0.116		0.300	
shjobw2	0.003	-0.003	0.012*	
	(0.004)	(0.005)	(0.006)	
	0.831	-0.618	2.045	
	0.406	0.537	0.041	
radhlw2	0.007	0.008	0.018*	0.010
	(0.005)	(0.004)	(0.007)	(0.006)
	1.381	1.813	2.555	1.576
	0.167	0.070	0.011	0.115
havmilsq		-0.000		
		(0.000)		
		-0.485		
		0.628		
shhlw2		0.006		
		(0.005)		
		1.218		
		0.223		
shrelaw2		-0.013*		
		(0.005)		
		-2.463		
		0.014		
suprtw2		0.001		
		(0.003)		
		0.241		
		0.809		
bf4				-0.083
				(0.050)
				-1.642
				0.101
bf40				-0.209*
				(0.102)
				-2.059
				0.039
Constant	-3.455***	-4.409***	-7.245***	-4.841***
	(0.728)	(0.789)	(1.207)	(1.283)
	-4.747	-5.586	-6.000	-3.774

	0.000	0.000	0.000	0.000
r2_p	0.191	0.211	0.366	0.300
bic	338.1512	396.5309	216.8723	248.9322
N	340	363	340	363

* p<0.05, ** p<0.01, *** p<0.001

```

66 .
67 . estout sexlife inthobbies vacations, cells(b(star fmt(%9.3f)) se(par) t p
> )          ///
>          stats(r2_p bic N, fmt(%9.3f %9.0g) )      ///
>          legend collabels(none) varlabels(_cons Constant)  ///
>          mlabels(sexlife intshobbies vacations) ///
>          title("main effects regression coefficients"
>          ///
>          "of reconstructed dose for females in wave 2")

```

main effects regression coefficients of reconstructed dose for females in wave
> 2

	sexlife	intshobbies	vacations
main			
age	0.063*** (0.016) 4.001 0.000	0.058*** (0.016) 3.496 0.000	0.056*** (0.017) 3.375 0.001
bf4	-0.129** (0.040) -3.241 0.001	-0.016 (0.050) -0.328 0.743	
illw2	0.249 (0.253) 0.985 0.325		
whppain	0.017 (0.011) 1.517 0.129	0.004 (0.010) 0.416 0.677	0.010 (0.009) 1.126 0.260
whpsleep	0.005 (0.007) 0.686 0.492	0.027*** (0.006) 4.255 0.000	0.012* (0.006) 2.041 0.041
whpel	0.007 (0.006) 1.130 0.258	0.009 (0.007) 1.349 0.177	0.007 (0.006) 1.209 0.227
avgcumdosew2	0.039	0.069	0.042

	(0.048)	(0.089)	(0.091)
	0.817	0.769	0.465
	0.414	0.442	0.642
radhlw2	0.010	0.010	0.017**
	(0.006)	(0.006)	(0.006)
	1.711	1.687	2.798
	0.087	0.092	0.005
bf30		0.001*	
		(0.000)	
		1.961	
		0.050	
shrelaw2		-0.017**	
		(0.006)	
		-2.638	
		0.008	
suprtw2		-0.013**	
		(0.004)	
		-2.950	
		0.003	
deaw2			0.133
			(0.180)
			0.742
			0.458
shjobw2			-0.008
			(0.004)
			-1.732
			0.083
bf7m			-0.000
			(0.000)
			-0.781
			0.435
havmil			-0.000
			(0.001)
			-0.115
			0.908
Constant	-4.555***	-5.974***	-6.120***
	(1.106)	(1.344)	(1.047)
	-4.117	-4.445	-5.848
	0.000	0.000	0.000
r2_p	0.349	0.382	0.275
bic	275.7083	277.448	307.6772
N	340	363	363

* p<0.05, ** p<0.01, *** p<0.001

68 .

69 .