Mixing AutoMetrics with MARS: A study of psychological sequelae of Chornobyl

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2 Introduction

The current study conducted a random population sampling of Ukrainian residents in the Kiev and Zhytomir oblasts of that country, with the aim of developing long-term models of human nuclear disaster risk. Living in relatively close proximity to the Chornobyl Nuclear station in Ukraine, residents were exposed to the largest industrial radiological accident to date in 1986. A survey methodology was used to assess the complex bio-psycho-social pathways that contribute to long-term population outcomes after a significant radiological event. Data collection was conducted from 2008-2011.

In the effort to understand the long-term burden of nuclear accident exposure on a general population, investigators primary foci of interest included: the populations reconstructed cumulative dose exposure to ¹³⁷ Caesium (radiation source term for the Chornobyl event), cognitive perception of risk to health and environment, mental health status (standardized instruments), medical diagnoses (ICD-9), psychosocial functioning, health behavior, reproductive patterns, nutritional practices, Chornobyl accident information sources, and social communication networks. These domains were assessed in the population for their current status, and retrospectively for three earlier time periods from 1986 to the time of the survey interview. The current presentation describes preliminary statistical exploration of covariates influencing: 1) health-related behaviors as measured by the Nottingham Health Profile, 2) general mental health dysfunction as measured by the Brief Symptom Inventory , and 3) Chornobyl-related

post traumatic stress syndrome as measured by the Revised Civilian PTSD Scale.

Research of protracted, low dose radiation exposures from nuclear plant accidents, show relatively low impacts on population health risk [22]. Conversely, population perceptions of risk from radiological events remain active up to several years post-event [1, 3, 5, 14]. Toxic accidents (radiological, chemical, biologic agents) appear to drive a unique spectrum of psychosocial and behavioral responses in affected populations. These responses are distinct from event-related physical injury. They include depression, anxiety, traumatic response, increased medical services utilization, phobic nutritional behavior and changes in reproductive patterns [2, 3, 5, 6, 7].

3 Nomenclature

Multivariate adaptive regression splines (MARS)

- MARS is a method of using regression splines called basis functions to model local slope and intercept changes in relationships.
- It is similar to a piecewise regression model in that it combines basis functions to approximate the global nonlinear form of the relationship being modeled.
- MARS endeavors to improve the fit of the model and can identify interactions among the variables selected. To avoid overfitting, it uses a backward stepwise algorithm to trim the model to what it deems the essentials.

4 Data collection and sampling

- A sampling of 702 participants
- Randomized selection of phone numbers
- household sample of the Kiev and Zhytomir oblasts (states) of Ukraine
- Informed and consenting respondents were interviewered
- Responses were entered in interviewers handheld computers,
- uploaded for storage to a website constructed for the study (Vovici Corporation, USA).
- Participants ranged in age from 28-84.

5 Research instruments and measures

- 5.1 Outcome measures
- 5.1.1 Physical health
 - ICD-9 Medical Diagnosis. Diagnostic information from the Ukraine Ministry of Health database of annual standardized dispensary exams.
 - Randomized selection of phone numbers
 - Nottingham Health Profile. Standardized scale of self-reported health and its impact on multidomain behavioral functioning [11]
 - Reliability/validity on Russian language form tested in pilot study [12].

5.2 Outcome measures

5.2.1 Mental health

- Brief Symptom Inventory [13] Standardized scale measuring patterns of psychological distress: depression, anxiety, somatization, obsessive-compulsiveness, hostility, paranoia, psychoticism, global distress, positive symptom score. Russian form pilot tested [12].
- Revised Civilian PTSD Scale. Russian version anchored in Chornobyl event and restandardized [14]. Assesses post traumatic stress and distress clusters related to Chornobyl. Reliabilities computed in pilot study of this sample [12].
- of Psychosocial/Health Behavior: factors associated with toxic accident behavioral outcomes and are integrated within the research questionnaire, including medical services utilization, phobic behavior, substance use , eating habits, abortions and contraceptive use.

- demographics,
- perception of radiation risk and nuclear attitudes, Chornobyl cognitions [1, 8],
- Chornobyl information sources, accident characteristics (distance, relocation, etc.),
- general hazards perception,
- negative life events and buffers,
- coping style [7].
- \bullet cumulative external radiation exposure to $^{137} {\rm Caesium}$ was estimated for each participant (see below)

6 Objectives

Primary research objectives

- To explore the psychological sequelae of the Chornobyl disaster with respect to its effect on the part Ukrainian residents in the area.
- To analyze dose mental health effect on survivors and residents of the general area.
- To analyze perceived risk of exposure effect on survivors and residents of the general vicinity of such a nuclear event.

Primary methodological objectives

- Although we use AutoMetrics for variable selection, we want to ascertain whether MARS can add value to this process.
- To answer two questions concerning the value added to our AutoMetrics analysis by MARS.

Does MARS help us explore dimensions of our data which we might have ignored?

Does MARS aid in providing a well-behaved model for our investigation of the subject of choice?

Could MARS help us capture the nonlinear relationships between variables?

Secondary objectives

- To use general-to-specific (GETS) approach to minimize omitted variable bias [10].
- To assure ourselves of not having overlooked important relevant variables.
- To be able to search, identify, and model nonlinear functional relationships as painlessly as possible (while avoiding the curses of multidimensionality and multi- collinearity.
- We had a large number of variables– some of our datasets have approximately 3000 variables but we only have 339 males and 363 females in it.
- We could have more variables than observations in some cases and this is a problem AutoMetrics was designed to solve.
- To optimize construct validity– capturing all of the relevant and important dimensions
- To discover what disadvantages this usage entailed, if any.

7 Historical highlights in the development of MARS

- Automatic Interaction Detection: AID was developed by Morgan and Sonquist in1963 in JASA to identify interaction terms by testing relative reduction in sse. But there was a bias in the variable selection as the splits contained less observations.
- CHi-square Automatic Interaction Detection (CHAID) proposed by Gordon Kass in his doctoral dissertation in 1980 with bonferroni corrections for multiple tests.
- CART used regression and classification (Classification and regression trees) were later developed (Brieman, Freedman, Olshen and Stone, 1984) to combine categorical and continuous variables in the same tree output.
- Hastie, T. and Tibshirani, R. (1990) develop the generalized additive model which combines parametric and nonparametric estimation.
- Jerry Freedman developed MARS in 1993 as a method of automatic regression models— using basis functions and interactions between them.

Our situation

- Large dasets with more than 2900 variables in them.
- Sample size 363 women and 339 men
- Most analysis wrt socio-medical research is performed with gender-specific models.
- We want to developing optimal models to explain psychological phenomena experienced and reported by residents in the Oblasts near Chornobyl
- We have variables that exhibit a latency period before exhibiting a relationship
- We have variables that exhibit a delay manifesting a relationship.
- We depend on the multipath misspecification testing of AutoMetrics to show guide our analysis.

8 MARS algorithm

8.1 How MARS works

MARs automates variable selection. It transforms variables into truncated regression splines to accommodate nonlinear functional forms. It checks for potential interactions with predictors. It drops irrelevant variables to minimize overfitting [23, 2].

MARS models nonlinear components with a variation of piecewise spline regression models. It uses a recursive partitioning algorithm for the regression problem, which emerged in a program for classification and regression trees, dubbed CART (Brieman, L., Friedman, J., Olshen, R. and Stone, C., 1984).

Originally MARS recursively partitioned the domain into disjoint subregions, within each of which it split the domain into two left and right offspring regions. The regression components would reside either within the left or the right daughter domain. Basis functions have the general form of

$$f(x) = \beta_0 + \sum_{j=1}^{S} h(m)B(m)$$
 (1)

where $B(x) = I(x \in R_j)$ with I being an indictor function with a value of 1 if x is a condition is true or 0, otherwise. The indicator function then becomes part of a product of a series of step functions, in turn indicating membership of the subregion, formulated as

$$H[\eta] = 1, if \ \eta \ge 0, \tag{2}$$
$$= 0, \ otherwise.$$

such that they could be incorporated as follows [19]

$$B_5(x) = H[x_1 - 0] * H[1 - x_2] * H[x_2 - 0] * H[1 - x_1], (3)$$

which would delineate region 5 as a unit square within the larger region R. But as Stevens and Lewis point out these functions were step functions in disjointed areas. The disjoint requirement precluded MARS from finding linear additive models. Friedman replaced the univariate step functions with truncated regression splines.

A regression spline:

- If x = 2 4 6 8 10 12 14 16 18 20
- Spline formulation $(x-5)_+ = 0, 0, 0, 0, 0, 12, 14, 16, 18, 20$
- This spline may appear to be a shortened hockey stick with a value of zero until the 6th value is reached, upon the spline contains the values of x.

He formulated these in pairs, consisting of a primary basis function and a mirror image of it as

$$C = \{ (X_j - k_t)_+, (k_t - X_j)_+ \}$$
(4)

where k_t = represented the knots (joints) connecting the regression splines. The plus indicates that only the positive portion is used, so that the otherwise negative values are coded as zeros. This usually gives the basis function of the left-hand component with the parenthesis of Equation 5 the appearance of a hockey-stick. In effect, these functions entail recentering the variable around a knot. The effect of the transformation can take a variable that has a hockey - stick- like shape and linearize it, as shown in Figure 1.

8.2 basis function generatioon

- a stepwise algorithm. MARS begins with a constant and then searches each variable,
- and for each variable, tests all possible knots.
- The variable-knot combination that reduces the sum of squared errors the most or increases the sum of squared errors the least is selected.

Basis function transformation

- basis functions can transform hockey-stick like variables into more linear variables.
- For each primary transformation a mirror image is generated as well. For example if variable x is to be transformed, the basis function transformation might be max(0, 30 x) and its mirror image would be max (0, x 30).

What kinds of variables are amenable to such transformations?

- delayed responses or
- variables containing threshold conditions for a response to appear
- One example is the relationship between PTSD and anxiety shown in the left panel Figure 1, where a nonlinear variable is linearized by a basis function transformation. Each basis function can handle one sharp bend in the function. If more than one slope change is identified, several basis functions or an interaction may be required for their modeling.

The linearization renders the variable amenable to linear regression analysis, whereas the nonlinear form might be overlooked in a variable selection process.

The resulting linear combination of basis functions is then used in lieu of a nonlinear regression analysis. These functions are added to the linear combination of basis functions that constitutes the MARS model. For example, when PTSD in wave three is run against average cumulative reconstructed external dose of the respondent to 137 CS for women in Figure 1 or men in Figure 2, we can

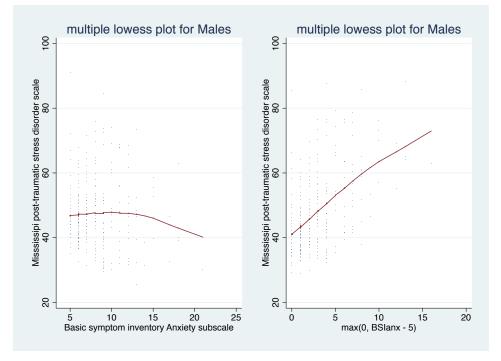


Figure 1: Comparison of PTSD anxiety relationship before and after basis function transformation

observe that the relationships are nonlinear.

MARS generates sets of pairs of truncated spline transformations for each possible position on the domain of the regressor. They are generated in reflected pairs, always representing the positive part of the domain. These plots often appear as hockey sticks of various dimensions. But just as our patterns of male PTSD against cumulative external dose of ¹³⁷Cesium reveal piecewise patterns, these hockey stick transformations can serve to linearize delayed, latent, or threshold responses to exogenous variables. For these reasons, relationships that have such a shape may be linearized by MARS and made more amenable to linear OLS regression analysis than otherwise might be the case. Because this is the first empirical examination of this subject, we are partly in an exploratory mode. Therefore, we want to be sure that we explore as many aspects of the relationships that might exist that we can. In situations such as these MARS may provide an invaluable investigative tool to plumb the depths of relationships that could exist within our dataset.

After generating many of these regression splines from the radial basis functions, the pair that diminished the sum of squared errors the most was the pair that was incorporated into the model.[17, 287].

To guard against overfitting by the forward stepwise algorithm, backward pruning was also incorporated into the system, while the model undergoes a 10-fold cross-validation test. Alternatively, the user may opt for a testing segment to be reserved, and then a training segment to preclude that overfitting.

The nonparametric selection of basis function provided a different driver for the model building process that could be modulated at a later stage. It is that last stage where AutoMetrics is brought into the process.

9 Dose reconstruction of exposure to ¹³⁷ CS

It is helpful to know that the ¹³⁷ CS in the Ukraine was not in general very large. Natural background radiation is about 2.4 mSv/year. Depending upon where people were and what they ate, and the extent to which they were outdoors, the amount of radiation to which they were ex-

posed may well have been below the level of normal biological reactivity.

However, that was not always the case. Infant children under the age of 2-5, when their thyroids absorb more iodine from the air, were more affected than older youth whose physical growth had slowed down. When natural iodine uptake from the air had considerably subsided on the part of the youth of five or older, as it does in the natural life cycle, the danger subsided. Moreover, the farther away they were from the accident site, the less their health was threatened, unless the winds shifted direction and carried the radioactive plume to them. For some time, there was considerable uncertainty as to who and how much the health of some people were threatened. In the Ukraine the in area we sampled, the situation can be described by Figure 2.

Because our sample was a random one, many of our respondents lived beyond the reach of the exclusion zone, an area of approximately 30 km from Chornobyl. Consequently, the mean effective dose sustained by this sample may not reflect the condition of those who were seriously and substantially exposed.

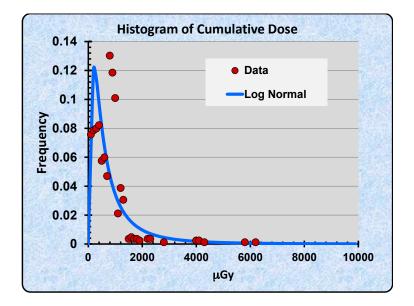


Figure 2: Cumulative $^{137}\mathrm{CS}$ dose in the Ukraine in microGrays

| Measure of ¹³⁷ CS | μ Grays |
|-------------------------------|-------------|
| minimum: | 44 |
| maximum: | 26,600 |
| mode: | 800 |
| median: | 715 |
| mean: | 838 |
| standard deviation: | 1695 |
| geometric mean: | 550 |
| geometric standard deviation: | 2.4 |

| Table 1: L | Deposition | of | 137 CS |
|------------|------------|----|-------------|
|------------|------------|----|-------------|

We hope to test this hypothesis with the best possible regression model. What does that mean? We have to have the optimal covariates in the model to assure that the general unrestricted model (GUM) is congruent. We wish to minimize the possibility of specification error by increasing the chance of including relevant covariates in our model. We use MARS to point out potential relationships that we might have missed by automatically and systematically finding relationships in the data by recentering and transforming our variables.

10 Application of MARS and AutoMetrics to variable selection and model building

10.1 BSI positive symptoms models

10.1.1 Female positive symptoms models

10.1.2 Female positive symptoms model using AutoMetrics alone

Table 2 Modelling BSIposymp by OLS-CS The dataset is: gals.in7 The estimation sample is: 1 - 363 Dropped 3 observation(s) with missing values from the sample

| _ | Coefficient | Std.Error | HACSE | t-HACSE | | Part.R ² |
|---------------|-------------------|---------------|---------|----------|--------|---------------------|
| Constant | 36.1449 | 7.460 | 6.939 | 5.21 | 0.0000 | |
| age | 0.567945 | 0.1656 | 0.1737 | 3.27 | 0.0012 | |
| edu3 | 8.27144 | 2.726 | 3.183 | 2.60 | 0.0098 | |
| edu4 | -1.86104 | 5.139 | 4.521 | -0.412 | 0.6809 | |
| marrw13 | 3.87829 | 3.807 | 4.073 | 0.952 | 0.3416 | |
| marrw16 | -4.68635 | 11.63 | 7.866 | -0.596 | 0.5517 | 0.0010 |
| marrw24 | 32.7286 | 23.94 | 6.326 | 5.17 | 0.0000 | 0.0732 |
| childw1 | -2.08219 | 2.009 | 2.003 | -1.04 | 0.2993 | 0.0032 |
| emplw16 | 10.1987 | 3.583 | 3.277 | 3.11 | 0.0020 | 0.0278 |
| emplw23 | 2.59385 | 4.633 | 4.253 | 0.610 | 0.5423 | 0.0011 |
| emplw24 | 2.61391 | 24.74 | 7.494 | 0.349 | 0.7275 | 0.0004 |
| depagw2 | 0.237815 | 0.07052 | 0.06819 | 3.49 | 0.0006 | 0.0346 |
| anxagw1 | 0.146863 | 0.03891 | 0.04042 | 3.63 | 0.0003 | |
| radfmw1 | 0.0553968 | 0.08792 | 0.07394 | 0.749 | 0.4542 | |
| radhlw1 | -0.180692 | 0.09119 | 0.08640 | -2.09 | 0.0372 | 0.0127 |
| radhlw3 | 0.234851 | 0.04998 | 0.05035 | 4.66 | 0.0000 | 0.0603 |
| inc4w1 | 9.02537 | 5.808 | 6.853 | 1.32 | 0.1887 | 0.0051 |
| inc1w3 | 10.2430 | 3.593 | 4.756 | 2.15 | 0.0320 | 0.0135 |
| avgcumdosew1 | U 4.72494 | 4.655 | 4.306 | 1.10 | 0.2733 | 0.0035 |
| avgcumdosew2 | U 4.50510 | 5.896 | 5.738 | 0.785 | 0.4329 | |
| avgcumdosew3 | U -3.55804 | 4.255 | 3.947 | -0.901 | 0.3680 | 0.0024 |
| | | | | | | |
| sigma | 23.0765 | RSS | | 0526.366 | | |
| R^2 | 0.365256 | | - | 0.000]** | | |
| Adj.R^2 | 0.327808 | 0 | | -1629.97 | | |
| no. of observ | | 1 1 1 | | 21 | | |
| mean(BSIposym | • | se(BSIposym | - | 28.1465 | | |
| | likelihood consta | | cluded: | | | |
| AIC | 6.33419 | SC | | 6.56088 | | |
| HQ | 6.42433 | FPE | | 563.590 | | |
| When the log- | likelihood consta | ant is includ | led: | | | |
| AIC | 9.17207 | SC | | 9.39876 | | |
| HQ | 9.26221 | FPE | | 9625.82 | | |
| Normality tes | t: Chi^2(2) = | 58.204 [0. | 0000]** | | | |
| Hetero test: | F(28, 329) = | - | - | | | |
| Hetero-X test | | 0.87766 [0. | - | | | |
| RESET23 test: | F(2,337) = | 1.9412 [0. | - | | | |
| | | | | | | |

10.1.3 Female model using AutoMetrics and MARS

Table 3 Modelling BSIposymp by OLS-CS The dataset is: gals.dta The estimation sample is: 1 - 363

| | | Coefficient | Std.Error | HACSE | t-HACSE | t-prob | Part.R ² |
|---------------|------|----------------|--------------|---------|----------|--------|---------------------|
| BFps2f | | 1.77871 | 0.1518 | 0.2158 | 8.24 | 0.0000 | 0.1629 |
| BFps3f | | 3.64609 | 0.07497 | 0.09200 | 39.6 | 0.0000 | 0.8182 |
| BFps3 | | -3.44030 | 0.1022 | 0.1302 | -26.4 | 0.0000 | 0.6667 |
| BFps4 | | 3.53564 | 0.07835 | 0.1000 | 35.3 | 0.0000 | 0.7816 |
| BFps8w2 | | -0.638613 | 0.1687 | 0.1754 | -3.64 | 0.0003 | 0.0366 |
| BFps9 | | -1.70877 | 0.09559 | 0.1195 | -14.3 | 0.0000 | 0.3692 |
| BFps10 | | 2.40125 | 0.1976 | 0.2177 | 11.0 | 0.0000 | 0.2585 |
| BFps11 | | -1.96492 | 0.1354 | 0.1634 | -12.0 | 0.0000 | 0.2930 |
| BFps12 | | 0.155755 | 0.04992 | 0.06037 | 2.58 | 0.0103 | 0.0187 |
| BFps22aw1 | | -0.346626 | 0.2059 | 0.1572 | -2.20 | 0.0281 | 0.0137 |
| BFps28w2 | | -4.59650 | 2.512 | 1.921 | -2.39 | 0.0173 | 0.0161 |
| avgcumdosew1 | U | 0.188333 | 1.400 | 1.189 | 0.158 | 0.8742 | 0.0001 |
| avgcumdosew2 | U | 1.86420 | 1.758 | 1.550 | 1.20 | 0.2300 | 0.0041 |
| avgcumdosew3 | U | -1.23718 | 1.263 | 1.287 | -0.961 | 0.3371 | 0.0026 |
| sigma | | 6.94119 | RSS | 1 | 6814.877 | | |
| R^2 | | 0.941297 | F(13, 349) = | 430.5 [| 0.000]** | | |
| Adj.R^2 | | 0.939111 | log-likelih | ood | -1211.24 | | |
| no. of observ | atio | ons 363 | no. of para | meters | 14 | | |
| mean(BSIposym | ıp) | 86.5069 | se(BSIposym | p) | 28.1296 | | |
| When the log- | like | elihood consta | nt is NOT in | cluded: | | | |
| AIC | | 3.91275 | SC | | 4.06295 | | |
| НQ | | 3.97245 | FPE | | 50.0384 | | |
| When the log- | like | elihood consta | nt is includ | ed: | | | |
| | | | ~~ | | | | |

6.90083

854.628B

```
// basis function legend
//------
// 0 bf4m = max(0, 32 - BSIsoma)
// 1 BFps2f = max(0, phobanx - 2.03628e-00*)
// 2 BFps3f = max(0, WHPer + 3.14933e-007)
// 3 BFps3 = max(0, WHPer - 33.7)
// 4 BFps4 = max(0, 33.7 - WHPer)
// 5 BFps8w2 = max(0, 5 - depagw2)
// 6 BFps9 = max(0, 5 - depagw2)
// 6 BFps9 = max(0, 3.852327e-007)
// 7 BFps10= max(0, BSIpsyc - 4)
// 8 BFps11 = max(0, 17- BSIpar)
// 9 BFps12 = max(0, MiPTSD-44)
// 10 BFps22aw1= max(0, 3-pillw1)
// 11 BFps28w2 = max(0, sepaw2 + 5.41039e-010)
Normality test: Chi^2(2) = 18.105 [0.0001]**
Hetero test: F(25,337) = 1.8159 [0.0108]*
Hetero-X test: F(97,265) = 2.0224 [0.0000]**
RESET23 test: F(2,347) = 1.5619 [0.2112]
```

6.75063 SC 6.81033 FPE

AIC HQ

10.2 Male Models for BSI Positive symptoms

Particular problems plague these men during wave two with significant effects. They include air and water pollution in both waves one and two (airw1 and airw2), and self-reported wave two depression and Chornobyl related health threat to oneself (radhlw2). However, the only dose-positive symptom impact is evident at wave one. The model as measured by adjusted $R^2=0.458$ is a respectable. But even more important is that only one of the model assumptions is in violation.

10.2.1 Male model with AutoMetrics Alone

Table 4: EQ(13) Modelling BSIposymp by OLS-CS The dataset is: guys.dta The estimation sample is: 1 - 340

| | | Coefficient | Std.Error | HACSE | t-HACSE | t-prob | Part.R^2 |
|--|------|----------------|--------------|---------|----------|---------|----------|
| Constant | | 76.0152 | 4.446 | 4.030 | 18.9 | 0.0000 | 0.5328 |
| educ4 | | 10.5398 | 4.830 | 5.188 | 2.03 | 0.0431 | 0.0131 |
| educ7 | | -13.1150 | 9.002 | 4.358 | -3.01 | 0.0028 | |
| marrw12 | | -15.9041 | 7.535 | 3.753 | -4.24 | 0.00020 | |
| marrw21 | | -12.1409 | 2.996 | 3.263 | -4.24 | | 0.0344 |
| marrw24 | | -19.2302 | 10.39 | 4.196 | -4.58 | 0.0002 | |
| marrw33 | | -6.72192 | 2.293 | 2.648 | -4.58 | 0.0000 | 0.0202 |
| childw2 | | -3.42953 | 1.581 | 1.863 | -1.84 | | 0.0202 |
| emplw25 | | -14.6481 | 4.954 | 4.564 | -3.21 | 0.0015 | 0.0320 |
| occ6w1 | | -14.7849 | 8.032 | 5.540 | -2.67 | 0.0010 | 0.0223 |
| occ8w1 | | -3.04011 | 2.796 | 3.050 | -0.997 | 0.3196 | 0.0032 |
| occ7w3 | | 9.52459 | 2.796 | 3.381 | 2.82 | 0.0052 | 0.0248 |
| inc1w1 | | 4.51683 | 3.031 | 2.836 | 1.59 | 0.1122 | 0.0081 |
| inc3w2 | | -2.83038 | 2.105 | 2.050 | -1.38 | 0.1684 | 0.0061 |
| inc4w2 | | -10.3680 | 6.052 | 5.953 | -1.74 | 0.0826 | 0.0096 |
| inc1w3 | | -1.23529 | 3.226 | 3.836 | -0.322 | 0.7476 | 0.0003 |
| inc4w3 | | 11.6291 | 5.642 | 4.657 | 2.50 | 0.0130 | 0.0196 |
| airw1 | | -0.0538877 | 0.03236 | 0.03894 | -1.38 | 0.1674 | 0.0061 |
| airw2 | | -0.108320 | 0.04605 | 0.04958 | -2.18 | 0.0296 | 0.0151 |
| airw3 | | 0.0562605 | 0.04006 | 0.04158 | 1.35 | 0.1770 | 0.0058 |
| radchw3 | | -0.0375368 | 0.03117 | 0.02909 | -1.29 | 0.1979 | 0.0053 |
| radh1w2 | | 0.224384 | 0.03387 | 0.04215 | 5.32 | 0.0000 | 0.0833 |
| illw3 | | 6.08420 | 1.103 | 1.337 | 4.55 | 0.0000 | 0.0623 |
| depagw1 | | 0.0908117 | 0.04081 | 0.06095 | 1.49 | 0.1372 | 0.0071 |
| depagw2 | | 0.269824 | 0.07069 | 0.1049 | 2.57 | 0.0106 | 0.0208 |
| avgcumdosew1 | U | 1.18106 | 1.284 | 0.5961 | 1.98 | 0.0484 | 0.0124 |
| avgcumdosew2 | U | -1.43189 | 4.240 | 2.289 | -0.626 | 0.5320 | 0.0013 |
| avgcumdosew3 | U | 0.560894 | 3.831 | 2.229 | 0.252 | 0.8015 | 0.0002 |
| | | | | | | | |
| sigma | | 17.0257 | RSS | | 0441.298 | | |
| R^2 | | 0.501136 | F(27, 312) = | | 0.000]** | | |
| Adj.R^2 | | 0.457965 | log-likelih | | -1431.64 | | |
| no. of observ | | | no. of para | | 28 | | |
| mean(BSIposym | | 74.9618 | se(BSIposym | | 23.1256 | | |
| 0 | lik | elihood consta | | cluded: | | | |
| AIC | | 5.74822 | SC | | 6.06354 | | |
| HQ | | 5.87386 | FPE | _ | 313.748 | | |
| | ·Lik | elihood consta | | ed: | 0 004 (0 | | |
| AIC | | 8.58609 | SC | | 8.90142 | | |
| HQ | | 8.71174 | FPE | | 5358.65 | | |
| Normality tes | st: | Chi^2(2) = | 35.773 [0. | **[0000 | | | |
| Hetero test: $F(39,300) = 1.3885 [0.0694]$ | | | | | | | |
| RESET23 test: | | F(2,310) = | | - | | | |
| | | - (2,020) | 1.0002. [0. | | | | |

| The da | | BSIposymp by t is guys.in7 | | | | | |
|---------------|------|-------------------------------|--------------|---------|----------|--------|---------------------|
| | | tion sample i | | | | | |
| | | Coefficient | | HACSE | t-HACSE | t-prob | Part.R [^] |
| depagw1 | | 0.0422038 | 0.01124 | 0.01496 | 2.82 | | |
| BFps1f | | 1.01014 | 0.3436 | 0.5369 | 1.88 | 0.0608 | 0.011 |
| BFps2f | | 3.77415 | 0.4887 | 1.001 | 3.77 | 0.0002 | 0.042 |
| BFps4 | | 0.800758 | 0.09437 | 0.1582 | 5.06 | 0.0000 | 0.074 |
| BFps14b | | 3.54876 | 0.5376 | 1.033 | 3.44 | 0.0007 | 0.035 |
| BFbsidep11 | | -0.0370855 | 0.01675 | 0.01997 | -1.86 | 0.0643 | 0.010 |
| BFbsidep12 | | 1.66140 | 0.1198 | 0.1356 | 12.3 | 0.0000 | 0.321 |
| BFbsidep13 | | 2.09106 | 0.1552 | 0.2106 | 9.93 | 0.0000 | 0.237 |
| BFbsidep14 | | 2.30043 | 0.1575 | 0.1948 | 11.8 | 0.0000 | 0.305 |
| BFptsd2a | | 0.0397467 | 0.01259 | 0.01901 | 2.09 | 0.0373 | 0.013 |
| BF5sociso3 | | -2.24763 | 0.3732 | 0.5600 | -4.01 | 0.0001 | 0.048 |
| BFdepx1 | | 0.776016 | 0.08013 | 0.1486 | 5.22 | 0.0000 | 0.079 |
| emplw11 | | 22.1433 | 4.671 | 10.71 | 2.07 | 0.0395 | 0.013 |
| emplw12 | | 16.3055 | 3.151 | 8.316 | 1.96 | 0.0508 | 0.012 |
| emplw13 | | 16.7090 | 3.211 | 8.284 | 2.02 | 0.0445 | 0.012 |
| emplw16 | | 16.9370 | 3.194 | 8.240 | 2.06 | 0.0407 | 0.013 |
| emplw22 | | 1.71709 | 0.6972 | 0.6371 | 2.70 | 0.0074 | 0.022 |
| occ5w2 | | 2.15340 | 1.128 | 1.186 | 1.82 | 0.0704 | 0.010 |
| occ4w3 | | 1.41793 | 0.8934 | 0.6351 | 2.23 | 0.0263 | 0.01 |
| I:180 | | -35.9714 | 5.501 | 3.896 | -9.23 | 0.0000 | 0.21 |
| avgcumdosew1 | U | -0.678928 | 0.3604 | 0.3830 | -1.77 | 0.0772 | 0.00 |
| avgcumdosew2 | U | 0.552981 | 1.172 | 1.380 | 0.401 | 0.6889 | 0.00 |
| avgcumdosew3 | U | -0.0585932 | 1.055 | 1.196 | -0.0490 | 0.9609 | 0.000 |
| sigma | | 4.85845 | RSS | 74 | 82.62731 | | |
| log-likelihoo | d | -1007.98 | | | | | |
| no. of observ | | ns 340 | no. of para | meters | 23 | | |
| mean(BSIposym | | 74.9618 | se(BSIposym | | 23.1256 | | |
| When the log- | | lihood consta | | | | | |
| AIC | | 3.22669 | SC | | 3.48570 | | |
| HQ | | 3.32989 | FPE | | 25.2013 | | |
| When the log- | like | lihood consta | nt is includ | ed: | | | |
| AIC | | 6.06456 | SC | | 6.32358 | | |
| HQ | | 6.16777 | FPE | | 430.424 | | |

10.2.2 Male model for Positive symptoms using AutoMetrics + MARS

// Basis function legend for positive symptom analysis //0 bf4m = max(0, 32 - BSIsoma) //1 BFps1f = max(0, bf4m - 1.19546e-007) //2 BFps2f = max(0, phobanx - 2.03628e-008) //3 BFps4 = max(0, 33.7 - WHPer) //4 BFbsidep11 = max(0, 65.06 - WHPsleep) //5 BFbsidep12 = max(0, BSIoc - 5) //6 BFbsidep13 = max(0, BSIanx - 5) //7 BFbsidep14 = max(0, BSIips - 4) //8 BFps14b = max(0, 11-BSIphanx) //10 BFptsd2a = max(0, fdferw2 - 7.17707e-007) //11 BF5sociso3 = max(0, 25-BSIsoma) //12 BFdepx1 = max(0, WHPer + 3.14933e-007) Normality test: Chi²(2) = 18.319 [0.0001]** Hetero test: F(37,301) = 6.5782 [0.0000]** RESET23 test: F(2,315) = 20.030 [0.0000]**

//-----

When we add basis functions to the model, we observe the results in Table 5. We find no evidence to support a dose-mental health relationship. We observe a tendency of the radial basis functions to dominate the output, with 47.8% of the variables selected are now basis functions. Because we now need variable labels to be able to facilitate an interpretation of the output, we provide a basis function index at the base of the model output.

The assumption violation rate increases again from 33.3% to 100%.

Many other factors appear to be related to general mental health on the part of the males, including depression in 1986 (depagw1 and BF5sociso3), somaticism (BFps1f), lack of sleep (BFbsidep11), anxiety (BFbsidep13 and BFbsi14b), interpersonal sensitivity (BFbsidep14), emotional reaction (BFps4 and BFdepx1), with some obsessive-compulsiveness (BSIdep12) included.

10.3 PTSD models

To test the hypothesis that cumulative dose predicts PTSD among men and women, we examine the addition of truncated regression splines does to the model of dose-PTSD relationship. First we want to graphically examine the relationship in Figure 3. Not only do we see that the relationship changes over time, but that at first it is characterized by an abrupt turn downward in wave one. In later waves, it more or less straightens out but retains a slight upward slope.

First we run our baseline model before adding the truncated spline transformations. Unlike our previous findings, we find that dose effect relationship seem to be associated with PTSD on the part of the male respondents. This association remains through all three waves.

We find comfort in believing that AutoMetrics has kept those covariates which were significantly related to the endogenous variable and which therefore provide us with the good control against specification error, at least insofar as the Ramsay reset test would have us believe. However, two of the regression assumptions are not fulfilled.

10.4 Male PTSD models

Socioeconomic status appears to play a significant role in PTSD for males. Being employed full-time during wave one (emplw12) is positively associated with PTSD. But part-time employment dur-

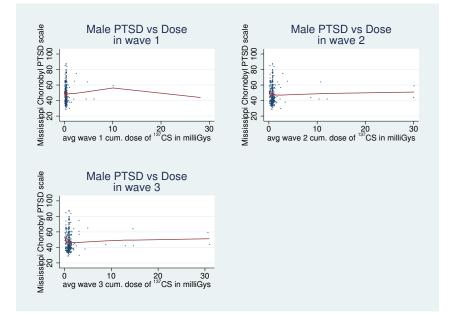


Figure 3: Male PTSD against Cumulative Effective Dose of 137 CS over three waves

ing wave three (emplw33) is significantly negatively associated with PTSD. For some reason, homemaking and caregiving becomes positively associated with male PTSD in wave three.

Many of these men have experienced some kind of catastrophe in wave three (cataw3) and they report illnesses in waves one and three (illw1 and illw3). There is a significant yet inverse relationship between their belief in the proportion of pollution due to Chornobyl (radchw3) in wave three. They are aware of the hazardous effects of radiation (efradw2) and remain fearful of eating radioactively contaminated food in wave three (fdferw3).

These are men who exhibit significant emotional reactions (WHPer), somaticism (BSIsoma), paranoia (BSIpar), and in wave two they report anxiety (anxagw2). For some reason the direction of significant dose-PTSD relationship switches from inverse to direct and then fades into relative statistical non-significance in wave three (p=0.074).

10.4.1 Male PTSD model using AutoMetrics alone

Table 6 EQ(20) Modelling MiPTSD by OLS-CS The dataset is:/guys.dta The estimation sample is: 1 - 340 Dropped 1 observation(s) with missing values from the sample

| | | Coefficient | Std.Error | HACSE | t-HACSE | +-nmah | Part.R^2 |
|--------------------|---|----------------|--------------|----------|----------|---------|----------|
| emplw12 | | 1.74095 | 0.8400 | 0.7437 | 2.34 | 0.0199 | |
| emp1w12 emp1w25 | | 9.57283 | 3.406 | 6.409 | 1.49 | 0.1363 | |
| emp1w23 emp1w33 | | -7.60494 | 1.677 | 1.891 | -4.02 | 0.0001 | |
| occ7w2 | | -8.96825 | 3.580 | 6.720 | -4.02 | 0.1830 | |
| occ7w2 | | 5.29819 | 1.689 | 1.802 | 2.94 | 0.0035 | |
| cataw3 | | 15.0846 | 3.670 | 5.033 | 3.00 | 0.0033 | 0.0200 |
| illw1 | | 2.82366 | 0.9845 | 1.149 | 2.46 | 0.0029 | 0.0270 |
| illw3 | | 1.56530 | 0.4444 | 0.4377 | 3.58 | 0.00143 | 0.0188 |
| shfamw1 | | -0.00984147 | 0.009452 | 0.009251 | -1.06 | 0.2882 | 0.0036 |
| suchrw3 | | 0.0144091 | 0.009432 | 0.009279 | 1.55 | 0.1214 | 0.0076 |
| fdferw3 | | 0.133797 | 0.01621 | 0.01669 | 8.02 | 0.0000 | 0.1690 |
| efradw2 | | 0.0607720 | 0.01242 | 0.01325 | 4.59 | 0.0000 | 0.0624 |
| radchw1 | | 0.0134653 | 0.01242 | 0.01344 | 1.00 | 0.3170 | 0.0032 |
| radchw3 | | -0.0395825 | 0.01282 | 0.01374 | -2.88 | 0.0042 | 0.0256 |
| WHPer | | 0.112901 | 0.03244 | 0.03768 | 3.00 | 0.0042 | 0.0276 |
| 1BSItotal | | 7.31530 | 0.4480 | 0.4154 | 17.6 | 0.0000 | |
| BSIsoma | | 0.502610 | 0.09646 | 0.1127 | 4.46 | 0.0000 | 0.0592 |
| BSIips | | -0.337637 | 0.2135 | 0.1954 | -1.73 | 0.0850 | |
| BSIpar | | 0.393373 | 0.1629 | 0.1462 | 2.69 | 0.0075 | |
| anxagw2 | | 0.0638708 | 0.02325 | 0.02468 | 2.59 | 0.0101 | |
| avgcumdosew1 | U | -1.03625 | 0.4535 | 0.3685 | -2.81 | | |
| avgcumdosew2 | U | 3.53115 | 1.577 | 1.736 | 2.03 | 0.0427 | 0.0129 |
| avgcumdosew3 | U | -2.80259 | 1.423 | 1.566 | -1.79 | 0.0744 | 0.0100 |
| 0 | | | | | | | |
| sigma | | 6.1492 | RSS | 1 | 1948.787 | | |
| log-likelihoo | d | -1084.84 | | | | | |
| no. of observ | ati | ons 339 | no. of para | ameters | 23 | | |
| mean(MiPTSD) | | 47.174 | se(MiPTSD) | | 11.9189 | | |
| When the log- | lik | elihood consta | nt is NOT in | ncluded: | | | |
| AIC | | 3.69808 | SC | | 3.95766 | | |
| HQ | | 3.80152 | FPE | | 40.3781 | | |
| When the log- | When the log-likelihood constant is included: | | | | | | |
| AIC | | 6.53596 | SC | | 6.79554 | | |
| HQ | | 6.63940 | FPE | | 689.636 | | |
| | | | | | | | |

| Normality test: | Chi^2(2) | = | 7.3566 | [0.0253]* |
|-----------------|-----------|---|--------|-----------|
| Hetero test: | F(40,298) | = | 1.6234 | [0.0133]* |
| RESET23 test: | F(2,314) | = | 1.0049 | [0.3673] |

10.4.2 Male PTSD model using AutoMetrics + MARS

By adding basis functions, the model becomes considerably less parsimonious. In fact, the number of parameters increases to 65, although there is a huge decline in the residual sums of squares from 11948.787 to 6783.70162. The Schwartz criterion is used as an arbiter, this actually rises from 3.958 to about 4.16 with the addition of the new variables. The only basis function that was added in this case was a linear combination of the dependent variable, which we subsequently deleted from the variable candidate pool.

10.5 Male PTSD model with AutoMetrics + MARS- cont'd

Table 7 EQ(25) Modelling MiPTSD by OLS-CS The dataset is: guys.dta The estimation sample is: 1 - 340 Dropped 9 observation(s) with missing values from the sample

| | Coefficient | Std.Error | HACSE | t-HACSE | - | Part.R ² |
|-----------|-------------|-----------|---------|---------|--------|---------------------|
| emplw12 | 7.69107 | 2.819 | 2.420 | 3.18 | 0.0017 | 0.0366 |
| occ7w2 | -5.04297 | 1.831 | 1.736 | -2.91 | 0.0040 | 0.0308 |
| occ7w3 | 10.3099 | 2.226 | 2.254 | 4.57 | 0.0000 | 0.0729 |
| cataw3 | 8.28971 | 4.078 | 2.440 | 3.40 | 0.0008 | 0.0416 |
| illw1 | 5.17788 | 0.9376 | 0.9342 | 5.54 | 0.0000 | 0.1035 |
| fdferw3 | 0.0302069 | 0.03236 | 0.03779 | 0.799 | 0.4248 | 0.0024 |
| efradw2 | 0.0238300 | 0.01675 | 0.01587 | 1.50 | 0.1343 | 0.0084 |
| radchw1 | 0.0169080 | 0.01422 | 0.01504 | 1.12 | 0.2619 | 0.0047 |
| radchw3 | -0.0567978 | 0.01438 | 0.01427 | -3.98 | 0.0001 | 0.0562 |
| lBSItotal | 4.20572 | 1.033 | 0.9615 | 4.37 | 0.0000 | 0.0671 |
| BSIips | -0.911879 | 0.2381 | 0.2334 | -3.91 | 0.0001 | 0.0543 |
| anxagw2 | 0.0497071 | 0.02337 | 0.02327 | 2.14 | 0.0336 | 0.0169 |
| emplw13 | 5.89476 | 2.982 | 2.564 | 2.30 | 0.0223 | 0.0195 |
| emplw16 | 6.08702 | 2.940 | 2.568 | 2.37 | 0.0185 | 0.0207 |
| occ1w1 | -2.59177 | 1.095 | 0.9997 | -2.59 | 0.0101 | 0.0246 |
| occ2w2 | -3.66632 | 1.256 | 1.134 | -3.23 | 0.0014 | 0.0378 |
| occ3w2 | -3.72039 | 1.623 | 1.487 | -2.50 | 0.0129 | 0.0230 |
| occ4w2 | -3.96241 | 1.373 | 1.072 | -3.70 | 0.0003 | 0.0489 |
| occ8w2 | -2.27052 | 1.193 | 1.002 | -2.27 | 0.0243 | 0.0189 |
| occ1w3 | 10.7594 | 2.171 | 2.255 | 4.77 | 0.0000 | 0.0788 |
| occ2w3 | 11.5444 | 2.323 | 2.244 | 5.14 | 0.0000 | 0.0905 |
| occ3w3 | 10.5016 | 2.460 | 2.383 | 4.41 | 0.0000 | 0.0680 |
| occ4w3 | 9.37193 | 2.393 | 2.131 | 4.40 | 0.0000 | 0.0678 |
| occ5w3 | 11.9117 | 2.347 | 2.450 | 4.86 | 0.0000 | 0.0816 |
| occ6w3 | 6.08384 | 3.680 | 5.175 | 1.18 | 0.2408 | 0.0052 |
| inc2w3 | -2.02479 | 0.9227 | 0.9335 | -2.17 | 0.0310 | 0.0174 |
| inc3w3 | -2.38019 | 0.9749 | 1.008 | -2.36 | 0.0189 | 0.0205 |
| cataw1 | -1.43900 | 0.9247 | 0.8904 | -1.62 | 0.1072 | 0.0097 |
| cataw2 | 7.34793 | 3.360 | 2.425 | 3.03 | 0.0027 | 0.0334 |
| dvcew3 | -2.54128 | 1.302 | 1.335 | -1.90 | 0.0581 | 0.0134 |
| illw2 | 1.16389 | 0.5908 | 0.5776 | 2.02 | 0.0449 | 0.0150 |
| movew2 | 4.36285 | 0.8182 | 0.6689 | 6.52 | 0.0000 | 0.1379 |
| shjobw1 | 0.0436225 | 0.01294 | 0.01375 | 3.17 | 0.0017 | 0.0365 |
| shjobw2 | -0.0396988 | 0.01321 | 0.01286 | -3.09 | 0.0022 | 0.0346 |
| | | | | | | |

| | Table 7 - | - continued. | | | | |
|-------------------|----------------|--------------|----------|----------|----------|-------------|
| | Coefficie | ent Std.Erro | r HAC | SE t-HAC | SE t-pro | ob Part.R^2 |
| | | | | | - | |
| | | | | | | |
| shhlw2 | 0.0223050 | 0.01419 | 0.01356 | 1.65 | 0.1011 | 0.0101 |
| shfincw2 | -0.0219229 | 0.01411 | 0.01400 | -1.57 | 0.1187 | 0.0091 |
| shfincw3 | 0.0136628 | 0.01356 | 0.01175 | 1.16 | 0.2459 | 0.0051 |
| shhousw1 | -0.0436495 | 0.01381 | 0.01440 | -3.03 | 0.0027 | 0.0334 |
| shhousw3 | 0.0403242 | 0.01213 | 0.01228 | 3.28 | 0.0012 | 0.0389 |
| shrelaw2 | 0.0375747 | 0.01576 | 0.01483 | 2.53 | 0.0119 | 0.0236 |
| shrelaw3 | -0.0429053 | 0.01358 | 0.01398 | -3.07 | 0.0024 | 0.0342 |
| suprtw3 | 0.00665292 | 0.008794 | 0.008390 | 0.793 | 0.4285 | 0.0024 |
| sufamw1 | 0.0500712 | 0.01839 | 0.02032 | 2.46 | 0.0144 | 0.0223 |
| suchrw2 | -0.0200943 | 0.008790 | 0.009018 | -2.23 | 0.0267 | 0.0183 |
| fdferw2 | 0.0793413 | 0.03196 | 0.04034 | 1.97 | 0.0502 | 0.0143 |
| kmacc | 0.0823621 | 0.01387 | 0.02051 | 4.02 | 0.0001 | 0.0572 |
| injothr | 1.59261 | 0.7767 | 0.8111 | 1.96 | 0.0506 | 0.0143 |
| kmwork | -0.0861868 | 0.01378 | 0.02045 | -4.21 | 0.0000 | 0.0626 |
| polprw3 | 0.0358676 | 0.01554 | 0.01554 | 2.31 | 0.0217 | 0.0196 |
| airw2 | -0.0413224 | 0.01281 | 0.01355 | -3.05 | 0.0025 | 0.0338 |
| radw1 | 0.0322737 | 0.01101 | 0.01106 | 2.92 | 0.0038 | 0.0310 |
| radhlw1 | -0.0385425 | 0.01545 | 0.01560 | -2.47 | 0.0141 | 0.0224 |
| radhlw3 | 0.0455350 | 0.01831 | 0.01870 | 2.44 | 0.0155 | 0.0218 |
| healthef | 0.0176544 | 0.01307 | 0.01555 | 1.14 | 0.2572 | 0.0048 |
| icdxcnt | -0.395851 | 0.2306 | 0.2331 | -1.70 | 0.0906 | 0.0107 |
| HP2work | 1.58433 | 0.8974 | 0.7985 | 1.98 | 0.0483 | 0.0146 |
| HP2hmcare | -2.48808 | 0.9763 | 0.8842 | -2.81 | 0.0053 | 0.0289 |
| HP2pbfhm | 1.92806 | 1.379 | 1.448 | 1.33 | 0.1842 | 0.0066 |
| BSIposymp | 0.382404 | 0.05583 | 0.05262 | 7.27 | 0.0000 | 0.1656 |
| BSIdep | -0.617556 | 0.1875 | 0.1736 | -3.56 | 0.0004 | 0.0454 |
| BSIphanx | -1.03931 | 0.2448 | 0.2323 | -4.47 | 0.0000 | 0.0700 |
| BSIhos | -0.425587 | 0.1974 | 0.1798 | -2.37 | 0.0186 | 0.0206 |
| avgcumdosew1 U | 0.730692 | 0.7812 | 0.5074 | 1.44 | 0.1510 | 0.0077 |
| avgcumdosew2 U | -3.98866 | 2.462 | 1.707 | -2.34 | 0.0202 | 0.0201 |
| avgcumdosew3 U | 2.47899 | 1.851 | 1.423 | 1.74 | 0.0827 | 0.0113 |
| | | | | | | |
| sigma | 5.05001 | RSS | 67 | 83.70162 | | |
| log-likelihood | -969.505 | | | | | |
| no. of observatio | ons 331 | no. of para | meters | 65 | | |
| mean(MiPTSD) | 47.0242 | se(MiPTSD) | | 11.7831 | | |
| When the log-like | elihood consta | nt is NOT in | cluded: | | | |
| AIC | 3.41291 | SC | | 4.15955 | | |
| HQ | 3.71070 | FPE | | 30.5107 | | |
| When the log-like | elihood consta | nt is includ | led: | | | |
| AIC | 6.25079 | SC | | 6.99743 | | |
| HQ | 6.54858 | FPE | | 521.107 | | |
| | | | | | | |
| Normality test: | Chi^2(2) = | 0.97273 [0. | - | | | |
| Hetero test: | F(104,226)= | 1.1353 [0. | - | | | |
| RESET23 test: | F(2,264) = | 12.918 [0. | 0000]** | | | |
| | | | | | | |

With this model, the hypothesis that dose predicts PTSD is not supported by the data for males at wave one and three The only apparent statistically significant association appears during wave two (avgcumdosew2, b = -3.989, se=2.462,t-prob=0.0202). That wave, covering the years of 1987 through 1996, was one of substantial economic privation and difficulty, during which other stresses and strains may have become more salient in the minds of these men.

What were the significant covariates? One group consisted of socio-demographic characteristics—such as being employed full, part, or unemployed in wave one (emplw12 emplw13 and emplw16, respectively) were somehow positive associated with Civilian PTSD in wave three.

Some aspects of socio-economic status appear to be related to PTSD in wave three. Having been a professional or administrator in wave one was negatively related to PTSD. In wave 2, employment in administrative support or technical sales (occ2w2), the protective services (occ3w2), precision production, mechanical, or crafts fields (occ3w2), or being a student (occ8w2) at the time were negatively associated with PTSD in wave three, as measured by the civilian PTSD scale score. By wave 3, such employment – along with employment as a factory laborer (occ5w3) or in agriculture, forestry, trapping, or logging (occ6w3) with the exception of being a student, were positively associated with PTSD. Income sufficiency with resources enough to meet basic necessities plus a little left over (inc3w3) is negatively associated with PTSD. It appears as if socio-economic adequacy may be important.

Environmental factors may have played a role as well. Males suffering from PTSD were concerned about the distance of their residence and workplace from the scene of the accident (kmacc and kmwork). By wave two, their concerns about air and water pollution (airw2) were inversely associated with PTSD. The farther the distance from the accident site the less the PTSD. They were concerned about the Chornobyl related threats to their own health (radhlw3) in wave three, whereas at first any concern they had was inversely associated with PTSD (radhlw1). They were people whose reported illnesses count in waves one (illw1) and two (illw2) were significantly associated with PTSD. Their fear of eating radioactively contaminated food was of borderline statistical significance in wave two (fdferw2 p=0.0502), which became statistically nonsignificant later.

PTSD appears to be positively related to having observed a catastrophic event in wave two or three (cataw2 or cataw3). A belief in the proportion of radioactively contaminated area is directly related to PTSD. Stresses and hassles from job related matters were in wave one directly associated with PTSD but switched to a significant inverse related relationship in wave two (shjobw1 and shjobw2). Stresses and hassles due to housing matters were inversely related in wave one but directly related in wave three (shhousw1 and shhousw3). Stresses and hassles relating to relationships went from significantly positive to significantly negative in waves two and three (shrelaw2 and shrelaw3).

Forms of support were occasionally related to PTSD in wave3. Family support was significant only in wave one (sufamw1). Chornobyl survivor support was negatively related to PTSD in wave two (suchrw2).

These were individuals who exhibited general mental health dysfunctionality (BSIposymp) whose health problems impacted their work (HP2work). Both of these traits were significantly positively related to PTSD in wave three. There was a significant positive association with PTSD between the log of total BSI. There was a borderline significance in the relationship of having injured others and PTSD (p=0.0506). Their reported anxiety in wave two was statistically significant with PTSD in wave three (anxagw2,b = .049, p = 0.034).

They also exhibited a number of significant inverse relationships between psychological symptomatology and PTSD. These included depression (BSIdep), phobic anxiety (BSIphanx), hostility (BSIhos), interpersonal sensitivity (BSIips), and the impact of health issues on home care cooking, and repairs (HP2hmcare). It is possible that the economic difficulties exacerbated the PTSD problems in wave two so that there appeared to be a dose-PTSD male relationship at that time.

10.6 Female PTSD models using AutoMetrics alone

Before turning to the female PTSD models, it behooves us to graph the relationships observed between PTSD and average cumulative dose over the three waves, shown in Figure 4.

Wave one appears to be less linear in functional form than the others, so we may have more need of basis functions there, but the other two relationships appear to be fairly linear and without the need of transformation.

We can examine the female baseline model ex-

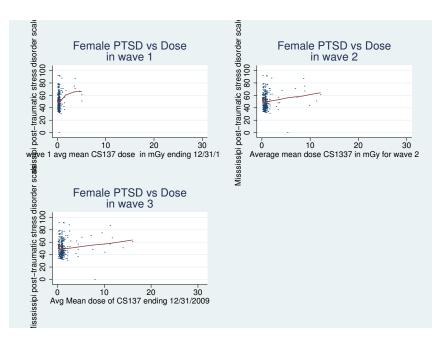


Figure 4: Female PTSD over three waves

plaining PTSD, for our hypothesis test of stressors, and then of buffers relating to the PTSD. In the baseline model, we observe no statistically significant relationship between cumulative external dose as measured by avgcumdose in waves one through three, and PTSD for women, measured by the Civilian PTSD scale score.

For the women, the baseline model has 55 parameters. Two-thirds of the first page of output consist in socio-economic attributes. How does socio-economic status relate to PTSD? Some of the largest coefficients indicate salient aspects of socio-economic status related to PTSD. Being single (marrw21) in wave two 2 can be positively related to it. Income insufficiency (inc1w3) or almost having such insufficiency (inc2w3) in wave 3 are significantly related to the PTSD. Having a child in wave one is significantly inversely related to it. With respect to occupational status, working outdoors (occ6w1) is significantly related to PTSD in wave two (farming, logging, trapping, etc).

Other stressors are salient. Accidents (accdw1) in wavelare significantly positively related to female PTSD in wave three. Stresses and hassles from relationships (shrelaw1) during wave one are significantly positively related to PTSD. Catastrophes in wave three are positive related to PTSD for women. General mental dysfunctionality as measured by the BSI positive symptom scale is significantly positively related to PTSD. Self-reported depression (depagw1 and depagw2) and anxiety (anxagw1 and anxagw2) may be statistically related as well. They seem to be significantly related but the signs of their parameter estimate change in opposite directions as we move from wave one to wave two. Health problem interference with interests and hobbies (HP2inthob) is significantly

positively associated with PTSD on the part of the female respondents. Consumption of pain pills (pillw2) in wave two seems to be statistically related to PTSD. The geodesic distance of the residence from the accident site (havmil) is also found to be statistically significantly related to PTSD. Also, the stresses and hassles from the job and from relationships appears to be significantly positively related to female PTSD.

Buffers to PTSD are also evident in this model. Having a child in wave one (childw1) for a woman is inversely related to PTSD. Being an MD (occ8w3) is also inversely related to PTSD. Working outdoors in agriculture, logging, forestry, are other jobs inversely related to PTSD (occ6w3). Health issues impacting one's sex life seem to be inversely related to PTSD for women (HP2sxlife). The stresses and hassles from one's health are inversely related to PTSD for women.

Supports also have a role to play. Family support(sufamw2) in wave two is significantly related as is Chornobyl survivor support in wave one (suchrw1). The belief that a large proportion of pollution is due to Chornobyl (radchw3) is significantly positively related to female PTSD.

| | stimation sample i | | | | | |
|----------|--------------------|----------------|----------|----------|--------|---------------------|
| Droppe | d 7 observation(s |) with missing | g values | from the | sample | |
| | Coefficient | Std.Error | HACSE | t-HACSE | t-prob | Part.R ² |
| marrw12 | 8.24237 | 4.931 | 4.806 | 1.72 | 0.0873 | 0.0096 |
| marrw21 | 12.3299 | 1.805 | 1.792 | 6.88 | 0.0000 | 0.1355 |
| marrw22 | 8.47335 | 3.703 | 2.766 | 3.06 | 0.0024 | 0.0301 |
| marrw23 | 6.23052 | 1.614 | 1.725 | 3.61 | 0.0004 | 0.0414 |
| childw1 | -1.26565 | 0.8791 | 0.7999 | -1.58 | 0.1146 | 0.0082 |
| childw2 | 2.55125 | 0.9794 | 0.8653 | 2.95 | 0.0034 | 0.0280 |
| emplw13 | 2.94446 | 2.047 | 1.774 | 1.66 | 0.0980 | 0.0090 |
| emplw32 | 3.52670 | 1.876 | 2.169 | 1.63 | 0.1050 | 0.0087 |
| emplw33 | 4.18612 | 7.971 | 1.777 | 2.36 | 0.0191 | 0.0180 |
| occ2w1 | 2.77946 | 1.582 | 1.613 | 1.72 | 0.0860 | 0.0097 |
| occ4w1 | -0.606786 | 2.170 | 1.844 | -0.329 | 0.7424 | 0.0004 |
| occ6w1 | 13.1978 | 4.550 | 2.815 | 4.69 | 0.0000 | 0.0678 |
| occ6w2 | -9.92894 | 4.473 | 2.773 | -3.58 | 0.0004 | 0.0407 |
| occ8w2 | -5.30301 | 2.045 | 2.187 | -2.42 | 0.0159 | 0.0191 |
| occ8w3 | -9.21187 | 8.387 | 2.808 | -3.28 | 0.0012 | 0.0344 |
| inc1w3 | 8.22857 | 1.680 | 1.812 | 4.54 | 0.0000 | 0.0639 |
| inc2w3 | 8.85438 | 1.391 | 1.337 | 6.62 | 0.0000 | 0.1267 |
| inc3w3 | 7.71199 | 1.413 | 1.372 | 5.62 | 0.0000 | 0.0947 |
| deaw1 | 1.16721 | 0.5685 | 0.6472 | 1.80 | 0.0723 | 0.0107 |
| deaw2 | -0.631489 | 0.5636 | 0.5662 | -1.12 | 0.2656 | 0.0041 |
| dvcew3 | -5.08077 | 1.884 | 1.790 | -2.84 | 0.0048 | 0.0260 |
| sepaw2 | -5.42709 | 3.268 | 2.621 | -2.07 | 0.0392 | 0.0140 |
| sepaw3 | 7.34660 | 2.257 | 2.257 | 3.26 | 0.0013 | 0.0339 |
| accdw1 | 8.14050 | 2.308 | 1.849 | 4.40 | 0.0000 | 0.0603 |
| accdw3 | 4.52347 | 1.349 | 1.181 | 3.83 | 0.0002 | 0.0463 |
| cataw1 | -2.97149 | 1.501 | 1.410 | -2.11 | 0.0359 | 0.0145 |
| cataw3 | 9.48043 | 5.937 | 5.320 | 1.78 | 0.0757 | 0.0104 |
| shjobw1 | 0.0459037 | 0.01897 | 0.02220 | 2.07 | 0.0395 | 0.0140 |
| shfamw2 | 0.0173869 | 0.02005 | 0.02067 | 0.841 | 0.4010 | 0.0023 |
| shhlw1 | -0.0632661 | 0.01947 | 0.02554 | -2.48 | 0.0138 | 0.0199 |
| shhlw3 | 0.0113486 | 0.01642 | 0.01671 | 0.679 | 0.4976 | 0.0015 |
| shfincw2 | -0.0151884 | 0.02226 | 0.02365 | -0.642 | 0.5212 | 0.0014 |
| shfincw3 | -0.0204790 | 0.01681 | 0.01948 | -1.05 | 0.2941 | 0.0036 |
| shrelaw1 | 0.0579751 | 0.01556 | 0.01523 | 3.81 | 0.0002 | 0.0458 |
| sufamw3 | 0.0354931 | 0.01329 | 0.01453 | 2.44 | 0.0151 | 0.0194 |
| suchrw1 | 0.0885848 | 0.04288 | 0.04238 | 2.09 | 0.0374 | 0.0143 |
| pillw2 | 0.246420 | 0.1558 | 0.1249 | 1.97 | 0.0493 | 0.0127 |
| pillw3 | -0.132645 | 0.09234 | 0.09077 | -1.46 | 0.1450 | 0.0070 |
| injselfr | 1.48063 | 1.100 | 1.147 | 1.29 | 0.1978 | 0.0055 |

Table 8 - continued...

| explanatory v | ar | Coefficient | Std.Error | HACSE | t-HACSE | t-prob | Part.R ² |
|---------------|------|---------------|--------------|----------|----------|--------|---------------------|
| radchw1 | | -0.00608758 | 0.01568 | 0.01719 | -0.354 | 0.7234 | 0.0004 |
| radchw3 | | 0.0455101 | 0.01787 | 0.01961 | 2.32 | 0.0210 | 0.0175 |
| WHPer | | 0.0124274 | 0.03620 | 0.04239 | 0.293 | 0.7696 | 0.0003 |
| HP2pbfhm | | -2.00175 | 1.695 | 1.999 | -1.00 | 0.3176 | 0.0033 |
| HP2sxlife | | -5.74679 | 1.371 | 1.523 | -3.77 | 0.0002 | 0.0450 |
| HP2inthob | | 7.84715 | 1.583 | 2.314 | 3.39 | 0.0008 | 0.0367 |
| BSIposymp | | 0.292820 | 0.02081 | 0.02386 | 12.3 | 0.0000 | 0.3328 |
| havmil | | 0.00610682 | 0.001950 | 0.002717 | 2.25 | 0.0253 | 0.0165 |
| depagw1 | | -0.0666916 | 0.02328 | 0.02283 | -2.92 | 0.0037 | 0.0275 |
| depagw2 | | 0.109470 | 0.03546 | 0.03460 | 3.16 | 0.0017 | 0.0321 |
| anxagw1 | | 0.0455276 | 0.01835 | 0.01819 | 2.50 | 0.0128 | 0.0203 |
| anxagw2 | | -0.0720609 | 0.02882 | 0.03571 | -2.02 | 0.0445 | 0.0133 |
| avgcumdosew1 | U | 3.03671 | 1.669 | 1.928 | 1.57 | 0.1164 | 0.0081 |
| avgcumdosew2 | U | 2.25711 | 2.073 | 2.156 | 1.05 | 0.2960 | 0.0036 |
| avgcumdosew3 | U | -2.93719 | 1.507 | 1.815 | -1.62 | 0.1067 | 0.0086 |
| sigma | | 7.70878 | RSS | 17 | 946.4376 | | |
| log-likelihoo | d | -1202.94 | | | | | |
| no. of observ | atio | ns 356 | no. of para | meters | 54 | | |
| mean(MiPTSD) | | 49.6208 | se(MiPTSD) | | 12.0886 | | |
| When the log- | like | lihood consta | nt is NOT in | cluded: | | | |
| AIC | | 4.22359 | SC | | 4.81136 | | |
| HQ | | 4.45739 | FPE | | 68.4392 | | |
| When the log- | like | lihood consta | nt is includ | led: | | | |
| AIC | | 7.06146 | | | 7.64923 | | |
| HQ | | 7.29527 | FPE | | 1168.91 | | |

| Normality test: | Cn1 2(2) | = | 4.9117 | [0.0050] |
|-----------------|-----------|---|--------|------------|
| Hetero test: | F(81,272) | = | 1.1388 | [0.2219] |
| RESET23 test: | F(2,300) | = | 37.383 | [0.0000]** |

When we mix AutoMetrics with MARS to test the female PTSD hypothesis, we get an elaborate rather than a parsimonious model. The hypothesis that dose is directly associated with PTSD measured by the Civilian Mississippi scale is consistent with parameter estimates for cumulative dose of ¹³⁷ CS in mGys (avgcumdosew1 and avgcumdosew3, both of which are statistically significant at the 0.05 level).

| Table 9 EQ(32) M The datas | odelling MiPTS et is:gals.dta | • | | | | |
|-------------------------------|----------------------------------|----------------|--------|----------|-----------|----------|
| | ation sample i | | | | | |
| | |) with missing | values | from the | sample | |
| Diopped o | 000011401011(0 | , | Turuos | | 2 amp 2 a | |
| | Coefficient | Std.Error | HACSE | t-HACSE | t-prob | Part.R^2 |
| marrw10 | 2.40743 | 3.127 | 2.103 | 1.14 | 0.2533 | 0.0044 |
| marrw12 | 8.70104 | 3.872 | 4.456 | 1.95 | 0.0518 | 0.0126 |
| marrw13 | 3.79369 | 1.166 | 1.149 | 3.30 | 0.0011 | 0.0351 |
| marrw21 | 3.82439 | 1.410 | 1.323 | 2.89 | 0.0041 | 0.0272 |
| marrw25 | 7.03501 | 3.217 | 2.954 | 2.38 | 0.0179 | 0.0186 |
| marrw32 | -4.37927 | 2.753 | 1.800 | -2.43 | 0.0155 | 0.0194 |
| marrw35 | -5.37582 | 1.928 | 2.308 | -2.33 | 0.0205 | 0.0178 |
| childw1 | -0.595462 | 0.5982 | 0.7546 | -0.789 | 0.4307 | 0.0021 |
| emplw13 | 3.04047 | 1.861 | 1.736 | 1.75 | 0.0810 | 0.0102 |
| emplw15 | 5.87467 | 3.790 | 2.599 | 2.26 | 0.0245 | 0.0168 |
| emplw22 | 2.45999 | 1.227 | 1.047 | 2.35 | 0.0194 | 0.0181 |
| emplw25 | -3.04349 | 2.065 | 2.118 | -1.44 | 0.1518 | 0.0069 |
| emplw33 | 8.96577 | 7.031 | 2.166 | 4.14 | 0.0000 | 0.0542 |
| occ3w1 | -5.70773 | 1.423 | 1.598 | -3.57 | 0.0004 | 0.0409 |
| occ6w1 | 5.11728 | 2.547 | 2.332 | 2.19 | 0.0290 | 0.0159 |
| occ2w2 | 2.73149 | 1.220 | 1.465 | 1.86 | 0.0632 | 0.0115 |
| occ7w2 | 5.13912 | 1.805 | 1.523 | 3.38 | 0.0008 | 0.0367 |
| occ8w2 | -3.99304 | 1.859 | 1.748 | -2.28 | 0.0231 | 0.0172 |
| occ8w3 | -7.08976 | 6.967 | 1.908 | -3.72 | 0.0002 | 0.0442 |
| inc1w3 | 2.59772 | 1.277 | 1.311 | 1.98 | 0.0485 | 0.0130 |
| inc2w3 | 2.52167 | 0.9265 | 0.9327 | 2.70 | 0.0073 | 0.0239 |
| deaw2 | -0.776765 | 0.5023 | 0.5270 | -1.47 | 0.1416 | 0.0072 |
| dvcew1 | 3.21334 | 5.055 | 2.772 | 1.16 | 0.2473 | 0.0045 |
| dvcew2 | -6.65500 | 2.146 | 1.942 | -3.43 | 0.0007 | 0.0378 |

Continued on the next page...

| | C | Coefficient | Std.Error | HACSE | t-HACSE | - | Part.R ² |
|---------------|---------|-------------|-------------|----------|----------|--------|---------------------|
| sepaw1 | | -11.6999 | 5.853 | 5.004 | -2.34 | 0.0200 | 0.0180 |
| sepaw3 | | 5.04579 | 1.756 | 1.970 | 2.56 | 0.0109 | 0.0215 |
| accdw1 | | 8.29547 | 1.976 | 2.055 | 4.04 | 0.0001 | 0.0517 |
| accdw3 | | 4.28951 | 1.165 | 1.373 | 3.12 | 0.0020 | 0.0316 |
| cataw3 | | 12.6758 | 5.373 | 5.723 | 2.21 | 0.0275 | 0.0161 |
| illw3 | | 1.44130 | 0.4337 | 0.4956 | 2.91 | 0.0039 | 0.0275 |
| movew3 | | 1.88537 | 2.146 | 1.747 | 1.08 | 0.2813 | 0.0039 |
| shjobw1 | | 0.0507772 | 0.01648 | 0.01829 | 2.78 | 0.0058 | 0.0251 |
| shjobw3 | | 0.00812849 | 0.01277 | 0.01383 | 0.588 | 0.5572 | 0.0012 |
| shhlw1 | | -0.0444529 | 0.01734 | 0.02059 | -2.16 | 0.0317 | 0.0153 |
| shfincw1 | | 0.0448605 | 0.01388 | 0.01257 | 3.57 | 0.0004 | 0.0408 |
| shfincw3 | | -0.0395446 | 0.01394 | 0.01599 | -2.47 | 0.0139 | 0.0201 |
| suprtw2 | | -0.0222859 | 0.009940 | 0.009618 | -2.32 | 0.0212 | 0.0176 |
| sufamw3 | | 0.00922673 | 0.01160 | 0.009733 | 0.948 | 0.3439 | 0.0030 |
| pillw3 | | -0.114727 | 0.06201 | 0.04760 | -2.41 | 0.0165 | 0.0191 |
| airw1 | | 0.0250637 | 0.01327 | 0.01270 | 1.97 | 0.0494 | 0.0129 |
| radchw1 | | -0.0600910 | 0.01499 | 0.01546 | -3.89 | 0.0001 | 0.0481 |
| radchw3 | | 0.0758247 | 0.01561 | 0.01648 | 4.60 | 0.0000 | 0.0661 |
| WHPsleep | | 0.0585702 | 0.01677 | 0.01846 | 3.17 | 0.0017 | 0.0326 |
| WHPsociso | | -0.0184077 | 0.02557 | 0.02705 | -0.680 | 0.4967 | 0.0015 |
| HP2sxlife | | -4.72120 | 1.209 | 1.364 | -3.46 | 0.0006 | 0.0385 |
| HP2vacatn | | 5.48797 | 1.353 | 1.482 | 3.70 | 0.0003 | 0.0439 |
| BSIposymp | | 0.208521 | 0.01950 | 0.02045 | 10.2 | 0.0000 | 0.2579 |
| depagw1 | | -0.0678127 | 0.02029 | 0.02263 | -3.00 | 0.0030 | 0.0292 |
| depagw2 | | 0.100628 | 0.03216 | 0.03028 | 3.32 | 0.0010 | 0.0356 |
| anxagw1 | | 0.0612168 | 0.01615 | 0.01665 | 3.68 | 0.0003 | 0.0432 |
| anxagw2 | | -0.157812 | 0.03253 | 0.03638 | -4.34 | 0.0000 | 0.0592 |
| BFptsd5 | | 0.266666 | 0.04240 | 0.03959 | 6.74 | 0.0000 | 0.1318 |
| BFptsd2a | | 0.277405 | 0.03833 | 0.03685 | 7.53 | 0.0000 | 0.1593 |
| BFptsd4a | | 0.0842508 | 0.04848 | 0.04852 | 1.74 | 0.0835 | 0.0100 |
| BFptsd5a | | -0.0270958 | 0.01449 | 0.01617 | -1.68 | 0.0949 | 0.0093 |
| BFptsdw33 | | 0.0974791 | 0.03504 | 0.03758 | 2.59 | 0.0100 | 0.0220 |
| avgcumdosew1 | U | 4.28430 | 1.459 | 1.855 | 2.31 | 0.0216 | 0.0175 |
| avgcumdosew2 | U | 2.65230 | 1.847 | 1.800 | 1.47 | 0.1417 | 0.0072 |
| avgcumdosew3 | U | -3.54964 | 1.341 | 1.492 | -2.38 | 0.0180 | 0.0186 |
| | | | | | | | |
| sigma | | 6.68318 | RSS | 13 | 354.8046 | | |
| log-likelihoo | | -1155.8 | _ | | | | |
| no. of observ | vations | | no. of para | ameters | 59 | | |
| mean(MiPTSD) | | 49.6117 | se(MiPTSD) | | 12.0553 | | |
| When the log- | likeli | | | ncluded: | | | |
| AIC | | 3.94871 | SC | | 4.58824 | | |
| HQ | | 4.20305 | FPE | | 52.0259 | | |
| When the log- | ∙⊥ikeli | | | ded: | | | |
| AIC | | 6.78658 | SC | | 7.42611 | | |
| HQ | | 7.04093 | FPE | | 888.574 | | |
| | | | | | | | |

Table 9 continued from previous page ...

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Normality test:Chi^2(2) =5.0543 [0.0799]Hetero test:F(88,267) =1.4085 [0.0201]*RESET23 test:F(2,297) =1.4425 [0.2380]

Basis function legend: BFptsd5 = max(0, 70-fdferw2) BFptsd4a = max(0, 40 - depagw3) BFptsd5a = max (0, airw2-10) BFpstdw33 = max(0, anxagw3)

To summarize the female analysis, we will note some salient significant stressors and buffers for PTSD among women in several domains—that of socio- economic status (SES), that of major negative life events, that of daily stresses and hassles, environmental factors, and psychological sequelae. Using the partial R^2 as a form of beta weight, we see that within the domain of SES, being married in wave 1 (marrw13), being divorced in wave 2 (marrw25), unemployed in wave 3 (emplw35), and having a Ph.D. in wave 2(occ7w2) are major stressors relating to PTSD. Other SES stressors are income insufficiency for basic needs or borderline income sufficiency for basic needs in wave three (inc1w3 and inc2w3).

Prominent socio-economic status buffers are being retired in waves 2 and 3 (emplw25 and emplw35) and being a medical doctor in waves two and three (occ8w2 and occ8w3).

Among the major negative life events, several salient stressors appear. Some of the worst are getting separated in wave one (sepaw1) and getting a divorce in wave two, that time of economic tribulation (dvcew2). However, accidents in wave one (accdw1) and three (accdw3) also emerge as substantial stressors.

Within the domain of daily stresses and hassles, financial ones in wave 1 are significantly positively related to PTSD. The reception of Chornobyl support in wave 2 and the consumption of pain pills in wave three (pillw3) not surprisingly appear to work as buffers.

As for environmental or contextual effects, in wave 1 the percent belief air and water pollution (airw1) is dangerous is related to PTSD, as is the proportion of pollution due to Chornobyl is high in wave 3 (radchw3) appears to be significantly negatively related to PTSD, although there was a belief that the air and water pollution was dangerous in wave 1 (airw1) and wave 2 (BFptsd5a), whereas by wave three the direction of this relationship reversed itself, so that it became a significantly positive in wave 3.

Fears of eating contaminated food appear to be significantly positively related to PTSD in wave 2 (BFptsd5).

Psychological symptoms appeared in depression and anxiety in waves one and two (depagw1 and depage2), where they went from inverse to direct over those two waves. However, anxiety went from positive to negative as people became used to the situation from wave one to wave two (anxagw1 and anxagw2). The basis functions reveal that depression and anxiety in wave three were related to PTSD as well.

These effects impacted the sleep, sex life, and vacation plans of the female respondents such that even general mental dysfunction became significantly positively related to PTSD. When the nature of the situation is considered, these women may not seem so unreasonable, after all.

11 Conclusion

11.1 Exploration of new dimensions

Does MARS help us explore dimensions we might otherwise have ignored? Sometimes it does and sometimes it does not. For the men, the proportion of variance explained tends to increase, but for the females it does not. In our study of the BSI positive symptom response, males exhibit an increase in the number of factors extracted, but females do not. Females, to the contrary, exhibit a slight decline in the number of factors extracted. In our PTSD models, the same pattern appears. The result may vary with the nature of the data.

11.2 Are MARS models well-haved?

MARS models tend to violate more assumptions than AutoMetrics built models and MARS basis function generation supervised by AutoMetrics does not guarantee well-behaved residuals. Sometimes MARS generated basis functions if selected by AutoMetrics proliferate assumption violations and sometimes they do not. The results at this juncture are not conclusive, although there is tentative evidence that MARS is motivated by optimizing fit and not fulfilling regression assumptions. Therefore the user must be careful about his use of MARS.

MARS interaction generation fails to include all component parts of the interaction and actually may prune some of them out if it improves the overall fit. The resulting model may be misspecified according to conventional statistical theory.

11.3 Can MARS capture nonlinear relationships ?

MARS can by adding basis functions approximate a nonlinear relationship between the dependent variable and the other variables in the model. However, it may try to do this by building one of its incomplete and unbalanced interaction vectors. This may not be what the user wants so he must be careful about such developments.

11.4 Content validity Tables

Does MARS improve content validty? The answer is sometimes and not always. If the dimensionality of the explanatory variables increases, then the content validity increases. But the use of MARS may actually decrease the dimensionality of the model, as shown among women with respect to their Positive symptom scale scores and their Chornobyl PTSD scores in Tables 2 and 3 below. For male respondents the application of MARS appeared to increase dimnsionality but did not necessarily improve fit.

Did MARS always improve the fit? If we examine in table 2 immediately above the proportion of common variance explained we find that among male BSI positive symptoms, it did. However, among females, the proportion of common variance explained declined. As for the PTSD, the application of MARS as a front end generator of Table 2: Construct validity assessment of BSI positive symptom models

| | AutoMetrics alone | MARS + AutoMetrics |
|--|-------------------|--------------------|
| Male | | |
| model | | |
| nfactors | 4 | 6 |
| Number of singletons | 0 | 0 |
| % common variance explained: | 87 | 94 |
| number of regression predictors | 28 | 23 |
| Number of regression assumption violations | 1 | 2 |
| Female | | |
| model | | |
| nfactors | 3 | 2 |
| Number of singletons | 0 | 0 |
| % common variance explained | 92 | 87 |
| Number of regression assumption violations | 0 | 0 |
| number of regression predictors | 21 | 14 |
| Number of regression assumption violations | 1 | 2 |

basis functions did improve proportion of variance explained among men but not in women. If the dimensionality of the predictor variable set actually decline slightly among women. Further research needs to be done to be sure, but if these few examples can serve as a guide, the value added by MARS is mostly to capture nonlinearities in relevant relationships.

| | AutoMetrics | MARS+AutoMetrics |
|--|-------------|------------------|
| Male | | |
| model | | |
| nfactors | 4 | 15 |
| Number of singletons | 0 | 0 |
| % common variance explained: | 85 | 88 |
| number of regression predictors | 23 | 65 |
| number of regression assumption violations | 1 | 3 |
| Female | | |
| model | | |
| nfactors | 10 | 9 |
| Number of singletons | 0 | 0 |
| % common variance explained | 86 | 84 |
| number of regression predictors | 54 | 59 |
| number of regression assumption violations | 1 | 1 |

Table 3: Construct validity assessment of PTSD models

11.5 Advantages of MARS

- automated
- fast

.

• builds additive model for you

11.6 Disadvantages of MARS

- adding them together to get a function
- no prima facie clear translation available sometimes
- you will need a translation table
- \bullet only emphasizes goodness of fit
- does not test regression assumptions
- interaction terms are not conventional

11.7 Caveats

- MARS may provide an optimal fit at the expense of violating the assumptions of regression analysis
- does not test regression assumptions
- interaction terms are not conventional in that they do not include the component terms conventionally required for proper interaction specification.

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